

Optimal design of experiments with very low average replication

R. A. Bailey



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LinStat, July 2012

Poster at LinStat 2010, Tomar, Portugal, July 2010:
'Unreplicated experiments in early stage breeding programs'
by Katarzyna Marczyńska and Stanisław Mejza.

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Single replication of new treatments; many check plots.
How to use the information from the check plots?

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Question from Brian Cullis and David Butler, Adelaide,
Australia, November 2010:

If we are not interested in control treatments, we know that we
can get more information by replacing the check plots by
second replicates of some treatments—can you prove this?

Abstract

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I shall compare designs under the A criterion when the average replication is much less than two.

Agricultural plant-breeding trials

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Example

There are 224 new varieties, with very little seed of each.

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How do you design the experiment?

Assume that

number of varieties $<$ number of plots

and

number of plots $\ll 2 \times$ (number of varieties).

$f(\omega)$ = variety on plot ω .

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$$Y_\omega = \tau_{f(\omega)} + \text{stuff depending on plots.}$$

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Assume that

$$Y_\omega = \tau_{f(\omega)} + \text{stuff depending on plots.}$$

We want to minimize

$$\sum_i \sum_{j \neq i} \text{Var}(\hat{\tau}_i - \hat{\tau}_j).$$

Simplest model

$$Y_{\omega} = \tau_{f(\omega)} + \epsilon_{\omega}$$

where

$$\mathbb{E}(\epsilon_{\omega}) = 0, \quad \text{Var}(\epsilon_{\omega}) = \sigma^2,$$

$$\text{and } \text{Cov}(\epsilon_{\omega}, \epsilon_{\omega'}) = 0 \quad \text{if } \omega \neq \omega'.$$

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The A-optimal design has
2 plots for some varieties and 1 plot for all other varieties,
and is completely randomized.

Simplest model: example

56 varieties have replication 2;
168 varieties have replication 1.

Simplest model: example

56 varieties have replication 2;
168 varieties have replication 1.

[illegible]

Simplest model: example

56 varieties have replication 2;
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[illegible]

A breeder says . . .

Unfair!

The single plot with my variety was in an infertile part of the field.

$$Y_{\omega} = \tau_{f(\omega)} + g(\omega) + \epsilon_{\omega}$$

where

$g(\omega)$ is a two-dimensional low-degree polynomial in ω ,

$$\mathbb{E}(\epsilon_{\omega}) = 0, \quad \text{Var}(\epsilon_{\omega}) = \sigma^2,$$

and $\text{Cov}(\epsilon_{\omega}, \epsilon_{\omega'}) = 0$ if $\omega \neq \omega'$.

Fixed spatial trend

$$Y_{\omega} = \tau_{f(\omega)} + g(\omega) + \epsilon_{\omega}$$

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Caliński, Mejza, Marczyńska ...:

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and several plots for a well-established but uninteresting
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use one plot for each new variety

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place the “control” plots in a grid;

use the “control” responses to estimate the polynomial trend;

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use one plot for each new variety

and several plots for a well-established but uninteresting
“control”;

place the “control” plots in a grid;

use the “control” responses to estimate the polynomial trend;

estimate each variety effect by subtracting the trend value from
its response.

Spatial trend: example

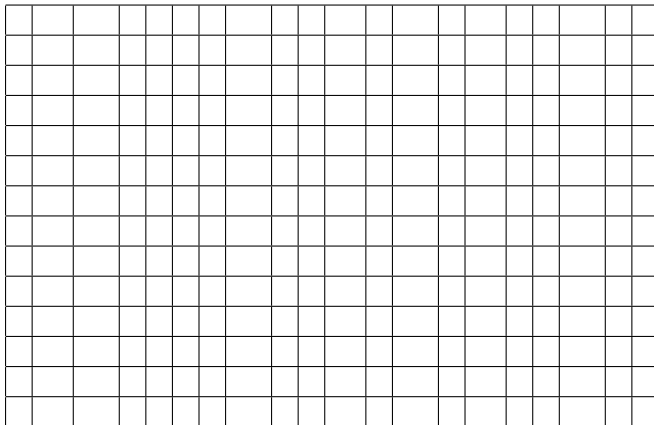
56 plots for “control”

224 new varieties have replication 1.

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		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		

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		X				X				X				X		
		X				X		3		X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
	2	X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X		1		X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		
		X				X				X				X		

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		X				X				X					X		
		X				X			3	X					X		
		X				X				X					X		
		X				X				X					X		
		X				X				X					X		
	2	X				X				X					X		
		X				X				X					X		
		X				X				X					X		
		X				X				X					X		
		X				X				X					X		
		X				X				X			1		X		
		X				X				X					X		
		X				X				X					X		
		X				X				X					X		
		X				X				X					X		

Controls are on every fifth plot, working along rows.

Spatial trend: example, another layout

56 plots for “control”

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Spatial trend: example, another layout

56 plots for “control”

224 new varieties have replication 1.

	X						X			X						X	
				X	X								X	X			
X								X	X								X
			X			X						X			X		
		X					X					X				X	
	X						X			X							X
				X	X								X	X			
X								X	X								X
			X			X						X			X		
		X					X					X				X	
	X						X			X							X
				X	X								X	X			
X								X	X								X
			X			X							X			X	

Controls are on every 5th plot, working boustrophedon along columns.

Spatial trend: example, a third layout

56 plots for “control”

224 new varieties have replication 1.

Spatial trend: example, a third layout

56 plots for “control”

224 new varieties have replication 1.

	X		X			X		X			X		X		
	X		X			X		X			X		X		
	X		X			X		X			X		X		
	X		X			X		X			X		X		
	X		X			X		X			X		X		
	X		X			X		X			X		X		
	X		X			X		X			X		X		

Controls are on a complete sub-rectangle.

Spatial trend: example, what should we optimize?

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X	X					X	X							X	X					X	X
X	X					X	X							X	X					X	X
X	X																			X	X
X	X					X								X						X	X
										X	X										
										X	X										
X	X					X								X						X	X
X	X																			X	X
X	X					X	X							X	X					X	X
X	X					X	X							X	X					X	X

Controls are positioned to make the **average** variance of prediction small if the trend is a polynomial of degree **three**.

Spatial trend: example, what should we optimize/assume?

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X	X	X						X	X	X	X						X	X	X
X	X	X															X	X	X
X																			X
X																			X
X								X	X										X
X								X	X										X
X								X	X										X
X								X	X										X
X																			X
X																			X
X	X	X															X	X	X
X	X	X						X	X	X	X						X	X	X

Controls are positioned to make the **maximum** variance of prediction small if the trend is a polynomial of degree **two**.

Minimizing the maximum variance of prediction is G-optimality.

Spatial trend: comments on optimality criteria

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Should we include the check plots in the average?

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Each criterion gives different designs for different assumptions about the trend.

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Thanks to Bradley Jones, who found these optimal designs.

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If the plots are not square, do the optimal designs change?

Yates (1936), Atiqullah and Cox (1962) consider controls spread throughout the field. In their analysis, a weighted mean of the response on the nearest controls is used as a covariate, rather than being simply subtracted.

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This is a rather different model.

Spatial correlation

where $Y_{\omega} = \tau_{f(\omega)} + \epsilon_{\omega}$

and $\mathbb{E}(\epsilon_{\omega}) = 0, \quad \text{Var}(\epsilon_{\omega}) = \sigma^2,$

$\text{Cov}(\epsilon_{\omega}, \epsilon_{\omega'})$ depends on the spatial relationship between ω and ω' .

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use one plot for each new variety and several plots for
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Kempton, Talbot, Besag, Martin, Eccleston . . . :
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place the “control” plots in some kind of grid;

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Kempton, Talbot, Besag, Martin, Eccleston . . . :
use one plot for each new variety and several plots for
“control”;
place the “control” plots in some kind of grid;
analyse all the data with GLS or REML.

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use 2 plots for some varieties and 1 plot for all other varieties,
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The field is partitioned into homogeneous blocks.
(One block has all the stony plots;
one block has all the plots near the trees;
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The field is partitioned into homogeneous blocks.
(One block has all the stony plots;
one block has all the plots near the trees;
one block has all the plots near the rabbit warren,)

$$Y_{\omega} = \tau_{f(\omega)} + \beta_{h(\omega)} + \epsilon_{\omega}$$

where

$$h(\omega) = \text{block containing } \omega,$$

$$\mathbb{E}(\epsilon_{\omega}) = 0, \quad \text{Var}(\epsilon_{\omega}) = \sigma^2,$$

$$\text{and } \text{Cov}(\epsilon_{\omega}, \epsilon_{\omega'}) = 0 \text{ if } \omega \neq \omega'.$$

Blocks: example

Rows are blocks, so there are 14 blocks, each with 20 plots.

Blocks: example, continued

224 varieties in 14 blocks of size 20.

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($280 - 224 = 56$ and $224 - 56 = 168$,

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Blocks: example, continued

224 varieties in 14 blocks of size 20.

($280 - 224 = 56$ and $224 - 56 = 168$,

so at least 168 varieties must have single replication.)

14 blocks	{	8 plots	12 plots	whole design Δ
		\vdots	\vdots	
		56 varieties	168 varieties all single replication	

Subdesign Γ has 56 varieties
in 14 blocks of size 8.

Blocks: remember that replication is very low

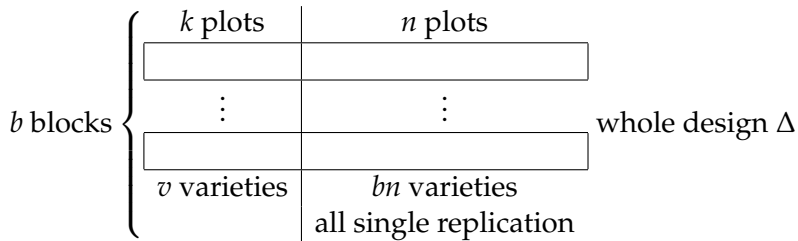
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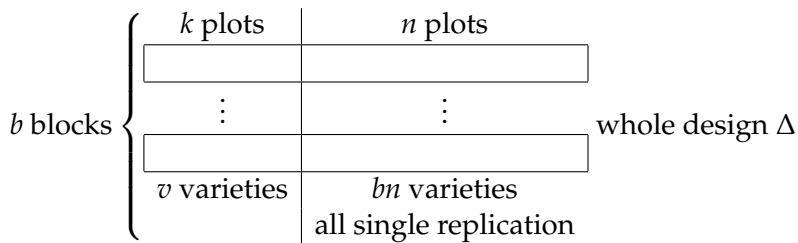
Yates (1936) concluded that square-lattice incomplete-block designs are superior to the inclusion of controls, but all of his examples had average replication equal to three or more.

Here we assume that average replication is (much) less than two.

A general block design with average replication less than 2

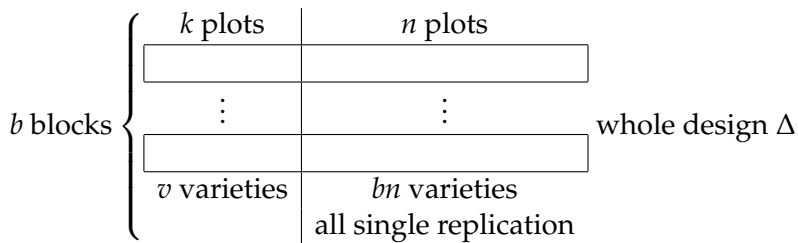


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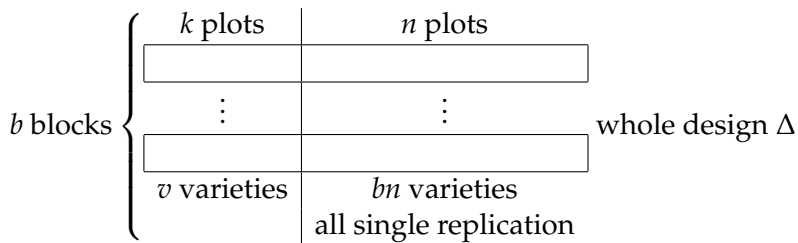
Whole design Δ has $v + bn$ varieties in b blocks of size $k + n$;

A general block design with average replication less than 2



Whole design Δ has $v + bn$ varieties in b blocks of size $k + n$;
the subdesign Γ has v **core** varieties in b blocks of size k ;

A general block design with average replication less than 2



Whole design Δ has $v + bn$ varieties in b blocks of size $k + n$;
the subdesign Γ has v **core** varieties in b blocks of size k ;
call the remaining varieties **orphans**.

Pairwise variance: two orphans in the same block

b blocks	{	k plots	n plots	whole design Δ
			$i \quad j$	
		\vdots	\vdots	
		v core varieties subdesign Γ	bn orphan varieties all single replication	

Pairwise variance: two orphans in the same block

b blocks	{	k plots		n plots	whole design Δ
				$i \quad j$	
		\vdots		\vdots	
		v core varieties subdesign Γ		bn orphan varieties all single replication	

$$\text{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = 2\sigma^2.$$

Pairwise variance: two orphans in different blocks

b blocks	{	k plots	n plots	whole design Δ
			i (block s)	
		\vdots	\vdots	
			j (block m)	
		v core varieties subdesign Γ	bn orphan varieties all single replication	

Pairwise variance: two orphans in different blocks

b blocks	{	k plots	n plots	whole design Δ
			i (block s)	
		\vdots	\vdots	
			j (block m)	
		v core varieties subdesign Γ	bn orphan varieties all single replication	

$$\text{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = 2\sigma^2 + \text{Var}_{\Gamma}(\hat{\beta}_s - \hat{\beta}_m).$$

Pairwise variance: two core varieties

b blocks	{	k plots	n plots	whole design Δ
		i		
		\vdots	\vdots	
		j		
		v core varieties subdesign Γ	bn orphan varieties all single replication	

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$$\text{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = \text{Var}_{\Gamma}(\hat{\tau}_i - \hat{\tau}_j).$$

Pairwise variance: one core variety and one orphan

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		i		
		\vdots	\vdots	
			j (block m)	
		v core varieties subdesign Γ	bn orphan varieties all single replication	

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b blocks	{	k plots	n plots	whole design Δ
		i		
		\vdots	\vdots	
			j (block m)	
		v core varieties subdesign Γ	bn orphan varieties all single replication	

$$\text{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = \sigma^2 + \text{Var}_{\Gamma}(\hat{\tau}_i + \hat{\beta}_m).$$

Sum of the pairwise variances

Theorem (cf Herzberg and Jarrett, 2007)

The sum of the variances of treatment differences in Δ

$$= \text{constant} + V_1 + nV_3 + n^2V_2,$$

where

V_1 = *the sum of the variances of treatment differences in Γ*

V_2 = *the sum of the variances of block differences in Γ*

V_3 = *the sum of the variances of sums of
one treatment and one block in Γ .*

(If Γ is equi-replicate then V_2 and V_3 are both increasing functions of V_1 .)

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(If Γ is equi-replicate then V_2 and V_3 are both increasing functions of V_1 .)

Consequence

For a given choice of k , make Γ as efficient as possible.

A less obvious consequence

Consequence

If n or b is large,
it may be best to make Γ a complete block design for k' controls,
even if there is no interest in comparisons between new
treatments and controls, or between controls.

$5n + 10$ treatments in 5 blocks of size $4 + n$

1	2	3	4	A_1	\cdots	A_n
---	---	---	---	-------	----------	-------

3	4	5	6	B_1	\cdots	B_n
---	---	---	---	-------	----------	-------

5	6	7	8	C_1	\cdots	C_n
---	---	---	---	-------	----------	-------

7	8	9	0	D_1	\cdots	D_n
---	---	---	---	-------	----------	-------

9	0	1	2	E_1	\cdots	E_n
---	---	---	---	-------	----------	-------

Youden and Connor (1953):
“experiments in physics do not need much replication because results are not very variable” —introduced chain block designs

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2	5	8	9	C_1	\cdots	C_n
---	---	---	---	-------	----------	-------

3	6	8	0	D_1	\cdots	D_n
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subdesign is dual of BIBD
(Herzberg and Andrews, 1978)

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subdesign is dual of BIBD,
best subdesign for $k = 4$

$5n + 10$ treatments in 5 blocks of size $4 + n$

1	2	3	4	A_1	\cdots	A_n
---	---	---	---	-------	----------	-------

1	5	6	7	B_1	\cdots	B_n
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2	5	8	9	C_1	\cdots	C_n
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subdesign is dual of BIBD,
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3	4	5	8	C_1	\cdots	C_n
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4	5	1	9	D_1	\cdots	D_n
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best subdesign for $k = 3$
is better for large n if $b \neq 5$

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best subdesign for $k = 3$
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K_1	K_2	1	2	A_1	\cdots	A_n
-------	-------	---	---	-------	----------	-------

K_1	K_2	3	4	B_1	\cdots	B_n
-------	-------	---	---	-------	----------	-------

K_1	K_2	5	6	C_1	\cdots	C_n
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K_1	K_2	7	8	D_1	\cdots	D_n
-------	-------	---	---	-------	----------	-------

K_1	K_2	9	0	E_1	\cdots	E_n
-------	-------	---	---	-------	----------	-------

better for large n if $b > 13$
even if there is no interest
in controls

The block design is D-optimal if it minimizes the volume of the ellipsoid of confidence for $(\tau_1, \dots, \tau_{v+bn})$ in the hyperplane defined by

$$\sum_i \tau_i = 0.$$

The **Levi graph** of the block design has

- ▶ one vertex for each treatment

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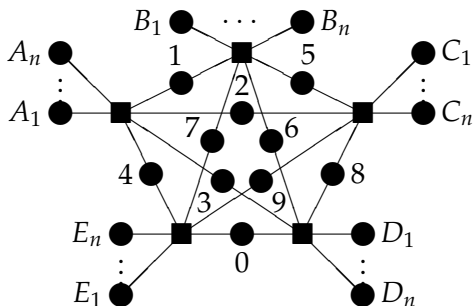
- ▶ one vertex for each treatment
- ▶ one vertex for each block

The **Levi graph** of the block design has

- ▶ one vertex for each treatment
- ▶ one vertex for each block
- ▶ one edge for each plot,
so that the edge corresponding to plot ω joins the vertices
corresponding to treatment $g(\omega)$ and block $h(\omega)$.

Levi graph: example

1	2	3	4	A_1	\dots	A_n
1	5	6	7	B_1	\dots	B_n
2	5	8	9	C_1	\dots	C_n
3	6	8	0	D_1	\dots	D_n
4	7	9	0	E_1	\dots	E_n

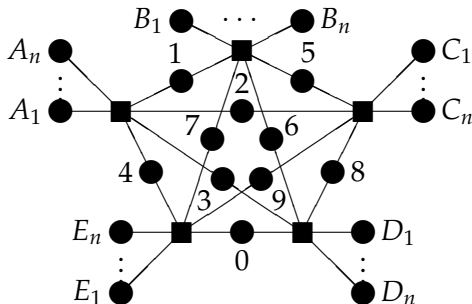


Spanning trees

A **spanning tree** for the Levi graph is a collection edges which provides a unique path between every pair of vertices.

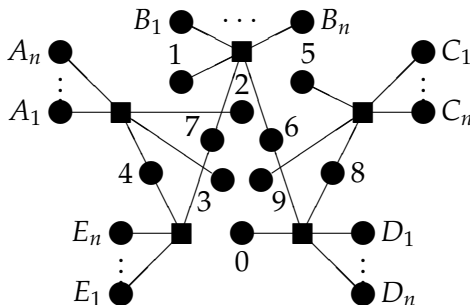
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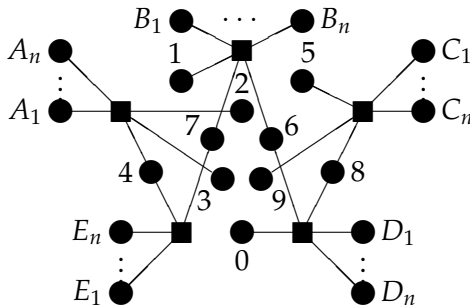
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Spanning trees

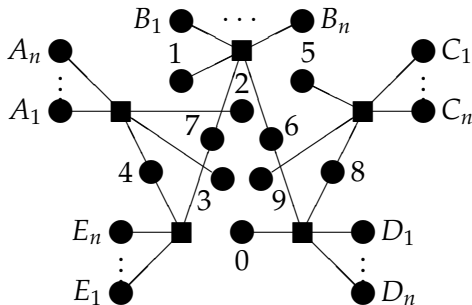
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This Levi graph has 8000 spanning trees; the Levi graph for the chain block design has 5120 spanning trees.

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The orphans make no difference to the number of spanning trees for the Levi graph.

Spanning trees and D-optimality

Theorem (Gaffke, 1981)

A block design is D-optimal if and only if its Levi graph maximizes the number of spanning trees.

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Consequence

The whole design Δ is D-optimal for $v + bn$ treatments in b blocks of size $k + n$ if and only if the core design Γ is D-optimal for v treatments in b blocks of size k .

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Consequence

Even when n or b is large, D-optimal designs do not include uninteresting controls.