Optimal design of experiments with very low average replication

R. A. Bailey

r.a.bailey@qmul.ac.uk

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I shall compare designs under the A criterion when the average replication is much less than two.

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Assume that

number of varieties < number of plots

and

number of plots $<< 2 \times$ (number of varieties).

 Y_{ω} = response on plot ω .

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We want to minimize

$$\sum_{i}\sum_{j\neq i}\operatorname{Var}(\hat{\tau}_{i}-\hat{\tau}_{j}).$$

$$Y_{\omega} = \tau_{f(\omega)} + \epsilon_{\omega}$$

where

$$E(\epsilon_{\omega} = 0), \quad \operatorname{Var}(\epsilon_{\omega}) = \sigma^{2},$$

and $\operatorname{Cov}(\epsilon_{\omega}, \epsilon_{\omega'}) = 0 \quad \text{if } \omega \neq \omega'.$

$$Y_{\omega} = \tau_{f(\omega)} + \epsilon_{\omega}$$

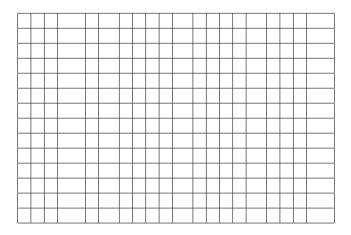
where

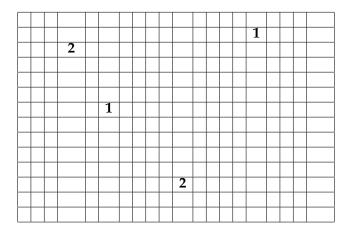
$$E(\epsilon_{\omega} = 0), \quad \operatorname{Var}(\epsilon_{\omega}) = \sigma^{2},$$

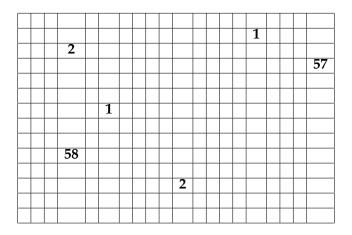
and $\operatorname{Cov}(\epsilon_{\omega}, \epsilon_{\omega'}) = 0 \quad \text{if } \omega \neq \omega'.$

The A-optimal design has

2 plots for some varieties and 1 plot for all other varieties, and is completely randomized.







Unfair!

The single plot with my variety was in an infertile part of the field.

$$Y_{\omega} = \tau_{f(\omega)} + g(\omega) + \epsilon_{\omega}$$

where

 $g(\omega)$ is a two-dimensional low-degree polynomial in ω ,

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Caliński, Mejza, ...: use one plot for each new variety and several plots for a well-established but uninteresting "control";

place the "control" plots in a grid;

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Caliński, Mejza, ...:

use one plot for each new variety

and several plots for a well-established but uninteresting "control";

place the "control" plots in a grid;

use the "control" responses to estimate the polynomial trend;

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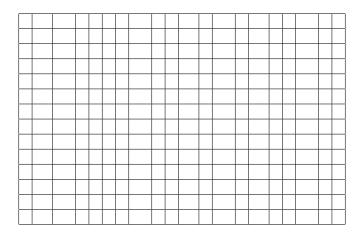
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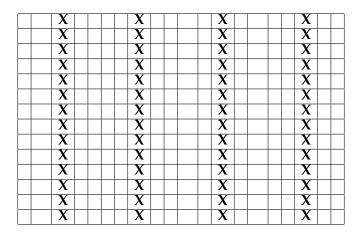
use one plot for each new variety

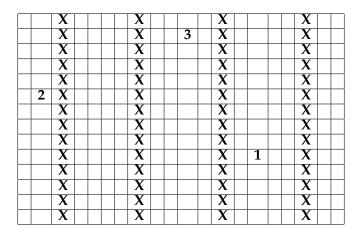
and several plots for a well-established but uninteresting "control";

place the "control" plots in a grid;

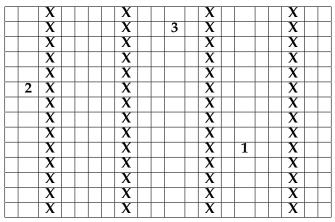
use the "control" responses to estimate the polynomial trend; estimate each variety effect by subtracting the trend value from its response.







56 plots for "control" 224 new varieties have replication 1.



Controls are on every fifth plot, working along rows.

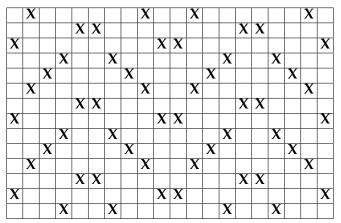
Spatial trend: example, another layout

56 plots for "control" 224 new varieties have replication 1.

Spatial trend: example, another layout

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Controls are on every 5th plot, working boustrophedon along columns.

Spatial trend: example, a third layout

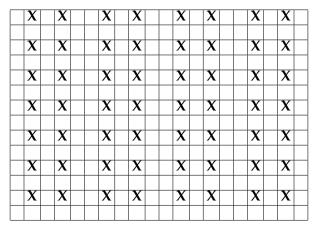
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Spatial trend: example, a third layout

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Controls are on a complete sub-rectangle

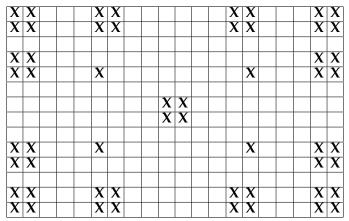
Spatial trend: example, what should we optimize?

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Controls are positioned to make the average variance of prediction small if the trend is a polynomial of degree three.

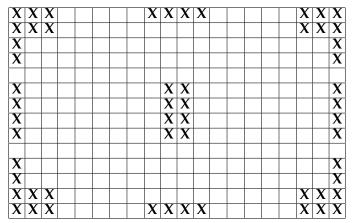
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Spatial trend: example, what should we optimize/assume?

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Controls are positioned to make the maximum variance of prediction small if the trend is a polynomial of degree two.

Yates (1936), Atiqullah and Cox (1962) consider controls spread throughout the field. In their analysis, a weighted mean of the response on the nearest controls is used as a covariate, rather than being simply subtracted.

where

$$Y_{\omega} = \tau_{f(\omega)} + \epsilon_{\omega}$$

$$E(\epsilon_{\omega}=0), \quad \operatorname{Var}(\epsilon_{\omega})=\sigma^{2},$$

and

 $\operatorname{Cov}(\epsilon_{\omega}, \epsilon_{\omega'})$ depends on the spatial relationship between ω and ω' .

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Kempton, Talbot, Besag, Martin, Eccleston ...: use one plot for each new variety and several plots for "control";

place the "control" plots in some kind of grid;

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Kempton, Talbot, Besag, Martin, Eccleston ...: use one plot for each new variety and several plots for "control";

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Blocks

The field is partitioned into homogeneous blocks. (One block has all the stony plots; one block has all the plots near the trees; one block has all the plots near the rabbit warren,) The field is partitioned into homogeneous blocks. (One block has all the stony plots; one block has all the plots near the trees; one block has all the plots near the rabbit warren,)

$$Y_{\omega} = \tau_{f(\omega)} + \beta_{h(\omega)} + \epsilon_{\omega}$$

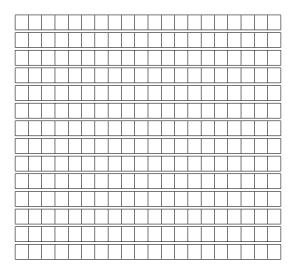
where

$$h(\omega) = \text{block containing } \omega,$$

$$E(\epsilon_{\omega} = 0), \quad \text{Var}(\epsilon_{\omega}) = \sigma^{2},$$

and
$$\text{Cov}(\epsilon_{\omega}, \epsilon_{\omega'}) = 0 \text{ if } \omega \neq \omega'.$$

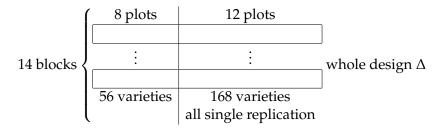
Rows are blocks, so there are 14 blocks, each with 20 plots.



Blocks: example, continued

224 varieties in 14 blocks of size 20.

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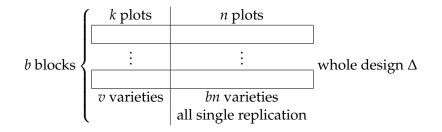


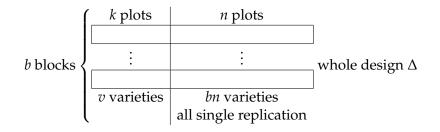
Subdesign Γ has 56 varieties in 14 blocks of size 8.

Yates (1936) concluded that square-lattice incomplete-block designs are superior to the inclusion of controls, but all of his examples had average replication equal to three or more.

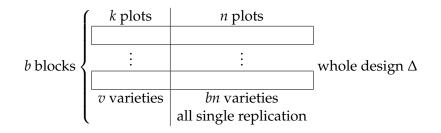
Yates (1936) concluded that square-lattice incomplete-block designs are superior to the inclusion of controls, but all of his examples had average replication equal to three or more.

Here we assume that average replication is (much) less than two.

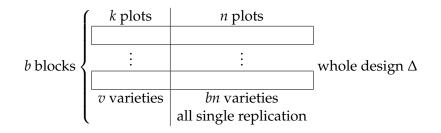




Whole design Δ has v + bn varieties in b blocks of size k + n;

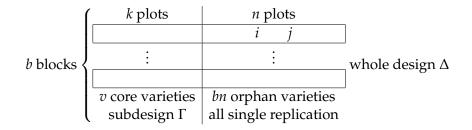


Whole design Δ has v + bn varieties in b blocks of size k + n; the subdesign Γ has v core varieties in b blocks of size k;

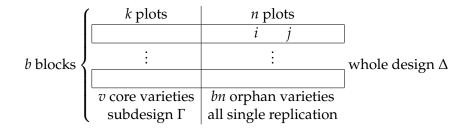


Whole design Δ has v + bn varieties in b blocks of size k + n; the subdesign Γ has v core varieties in b blocks of size k; call the remaining varieties orphans.

Pairwise variance: two orphans in the same block

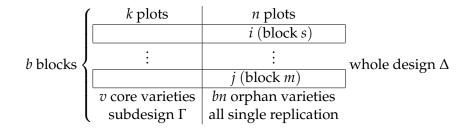


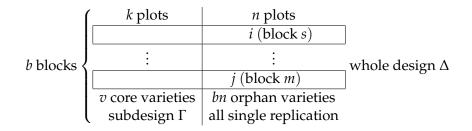
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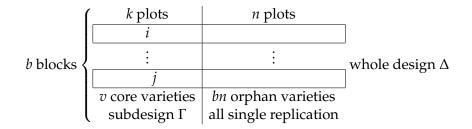
$$\operatorname{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = 2\sigma^2.$$

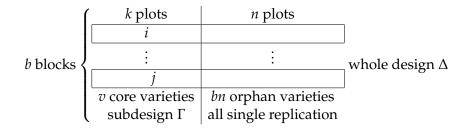
Pairwise variance: two orphans in different blocks





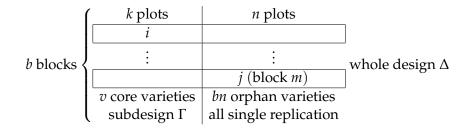
$$\operatorname{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = 2\sigma^2 + \operatorname{Var}_{\Gamma}(\hat{\beta}_s - \hat{\beta}_m).$$



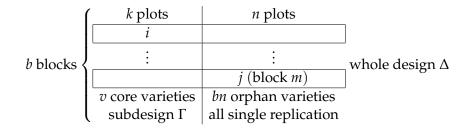


$$\operatorname{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = \operatorname{Var}_{\Gamma}(\hat{\tau}_i - \hat{\tau}_j).$$

Pairwise variance: one core variety and one orphan



Pairwise variance: one core variety and one orphan



$$\operatorname{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = \sigma^2 + \operatorname{Var}_{\Gamma}(\hat{\tau}_i + \hat{\beta}_m).$$

Sum of the pairwise variances

Theorem (cf Herzberg and Jarrett, 2007) The sum of the variances of treatment differences in Δ

$$= constant + V_1 + nV_3 + n^2V_2,$$

where

- V_1 = the sum of the variances of treatment differences in Γ
- $V_2 = the sum of the variances of block differences in \Gamma$
- V_3 = the sum of the variances of sums of one treatment and one block in Γ .

(If Γ is equi-replicate then V_2 and V_3 are both increasing functions of V_1 .)

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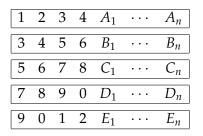
(If Γ is equi-replicate then V_2 and V_3 are both increasing functions of V_1 .)

Consequence

For a given choice of k, make Γ as efficient as possible.

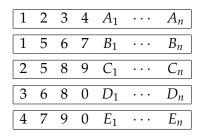
Consequence

If *n* or *b* is large, it may be best to make Γ a complete block design for *k*' controls, even if there is no interest in comparisons between new treatments and controls, or between controls.

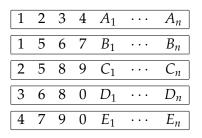


Youden and Connor (1953): "experiments in physics do not need much replication because results are not very variable" —introduced chain block designs

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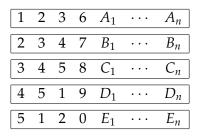
subdesign is dual of BIBD (Herzberg and Andrews, 1978)



subdesign is dual of BIBD, best subdesign for k = 4

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best subdesign for k = 3 is better for large *n* if $b \neq 5$



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$$\begin{bmatrix} K_1 & K_2 & 1 & 2 & A_1 & \cdots & A_n \\ \hline K_1 & K_2 & 3 & 4 & B_1 & \cdots & B_n \\ \hline K_1 & K_2 & 5 & 6 & C_1 & \cdots & C_n \\ \hline K_1 & K_2 & 7 & 8 & D_1 & \cdots & D_n \\ \hline \hline K_1 & K_2 & 9 & 0 & E_1 & \cdots & E_n \\ \end{bmatrix}$$

better for large *n* if b > 13 even if there is no interest in controls