# Design of experiments with very low average replication



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#### AGROSTAT 2012

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I shall compare designs under the A criterion when the average replication is much less than two.

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There are 224 new varieties, with very little seed of each. There are 280 plots available, in a  $14 \times 20$  rectangle. How do you design the experiment?

Assume that

number of varieties < number of plots

and

number of plots  $<< 2 \times$  (number of varieties).

 $f(\omega) =$  variety on plot  $\omega$ .

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 $\tau_i$  = effect of variety *i*.

Assume that

$$Y_{\omega} = \tau_{f(\omega)} + \text{stuff depending on plots.}$$

We want to minimize

$$\sum_{i}\sum_{j\neq i} \operatorname{Var}(\hat{\tau}_i - \hat{\tau}_j).$$

$$Y_{\omega} = \tau_{f(\omega)} + \epsilon_{\omega}$$

where

$$E(\epsilon_{\omega} = 0), \quad \operatorname{Var}(\epsilon_{\omega}) = \sigma^{2},$$
  
and  $\operatorname{Cov}(\epsilon_{\omega}, \epsilon_{\omega'}) = 0 \quad \text{if } \omega \neq \omega'.$ 

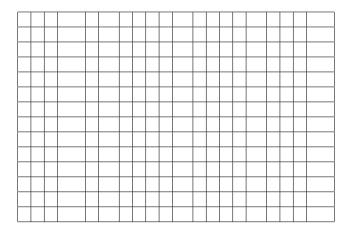
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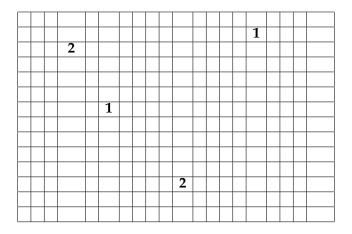
where

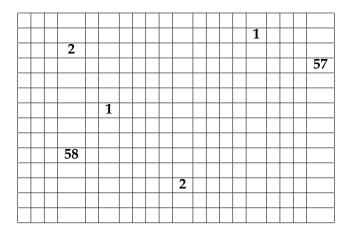
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and  $\operatorname{Cov}(\epsilon_{\omega}, \epsilon_{\omega'}) = 0 \quad \text{if } \omega \neq \omega'.$ 

The A-optimal design has

2 plots for some varieties and 1 plot for all other varieties, and is completely randomized.







#### Unfair!

The single plot with my variety was in an infertile part of the field.

$$Y_{\omega} = \tau_{f(\omega)} + g(\omega) + \epsilon_{\omega}$$

where

 $g(\omega)$  is a two-dimensional low-degree polynomial in  $\omega$ ,

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place the "control" plots in a grid;

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Caliński, Mejza, ...:

use one plot for each new variety

and several plots for a well-established but uninteresting "control";

place the "control" plots in a grid;

use the "control" responses to estimate the polynomial trend;

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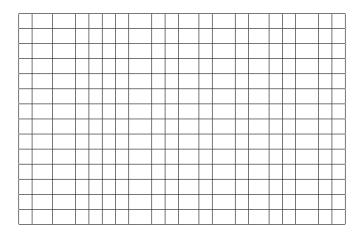
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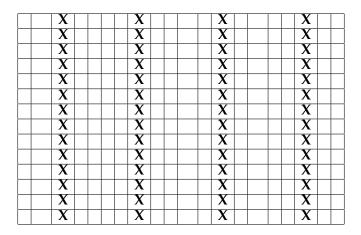
use one plot for each new variety

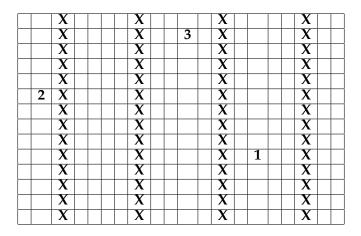
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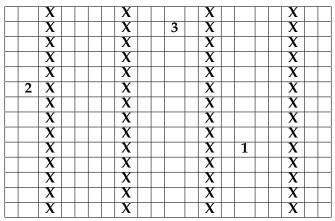
use the "control" responses to estimate the polynomial trend; estimate each variety effect by subtracting the trend value from its response.







#### 56 plots for "control" 224 new varieties have replication 1.



Controls are on every fifth plot, working along rows.

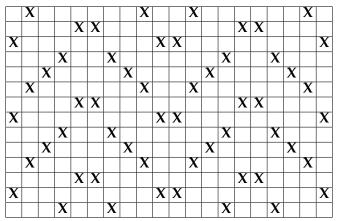
# Spatial trend: example, another layout

56 plots for "control" 224 new varieties have replication 1.

# Spatial trend: example, another layout

56 plots for "control"

224 new varieties have replication 1.



Controls are on every 5th plot, working boustrophedon along columns.

# Spatial trend: example, a third layout

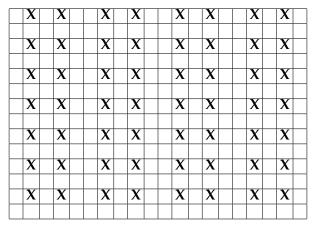
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## Spatial trend: example, a third layout

56 plots for "control"

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Controls are on a complete sub-rectangle

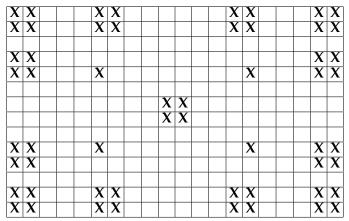
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Controls are positioned to make the average variance of prediction small if the trend is a polynomial of degree three.

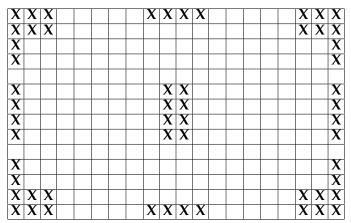
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Controls are positioned to make the maximum variance of prediction small if the trend is a polynomial of degree two.

Yates (1936), Atiqullah and Cox (1962) consider controls spread throughout the field. In their analysis, a weighted mean of the response on the nearest controls is used as a covariate, rather than being simply subtracted.

where

and

$$Y_{\omega} = \tau_{f(\omega)} + \epsilon_{\omega}$$

 $E(\epsilon_{\omega}=0), \quad \operatorname{Var}(\epsilon_{\omega})=\sigma^{2},$ 

 $Cov(\epsilon_{\omega}, \epsilon_{\omega'})$  depends on the spatial relationship between  $\omega$  and  $\omega'$ .

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Kempton, Talbot, Besag, Martin, Eccleston ...: use one plot for each new variety and several plots for "control";

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Kempton, Talbot, Besag, Martin, Eccleston ...: use one plot for each new variety and several plots for "control";

place the "control" plots in some kind of grid;

where

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place the "control" plots in some kind of grid; analyse all the data with GLS or REML.

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Cullis, Smith, Lim, Gilmour, Butler ...: use 2 plots for some varieties and 1 plot for all other varieties,

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Cullis, Smith, Lim, Gilmour, Butler ...: use 2 plots for some varieties and 1 plot for all other varieties, optimize the design by computer search,

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## **Blocks**

The field is partitioned into homogeneous blocks. (One block has all the stony plots; one block has all the plots near the trees; one block has all the plots near the rabbit warren, ....) The field is partitioned into homogeneous blocks. (One block has all the stony plots; one block has all the plots near the trees; one block has all the plots near the rabbit warren, ....)

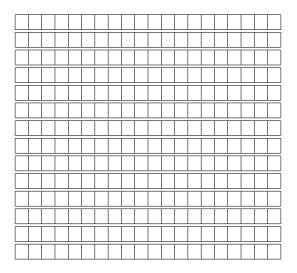
$$Y_{\omega} = \tau_{f(\omega)} + \beta_{h(\omega)} + \epsilon_{\omega}$$

where

$$h(\omega) = \text{block containing } \omega,$$
  

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and 
$$\text{Cov}(\epsilon_{\omega}, \epsilon_{\omega'}) = 0 \text{ if } \omega \neq \omega'.$$

Rows are blocks, so there are 14 blocks, each with 20 plots.



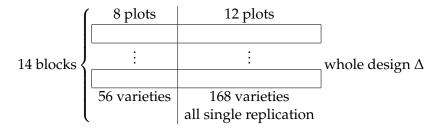
# Blocks: example, continued

224 varieties in 14 blocks of size 20.

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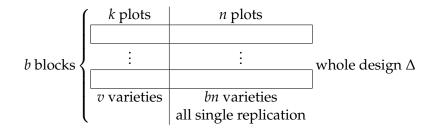
Subdesign  $\Gamma$  has 56 varieties in 14 blocks of size 8.

Yates (1936) concluded that square-lattice incomplete-block designs are superior to the inclusion of controls, but all of his examples had average replication equal to three or more.

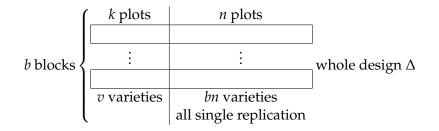
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Here we assume that average replication is (much) less than two.

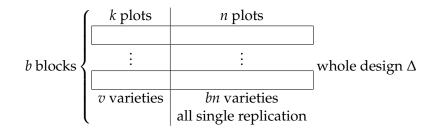
A general block design with average replication less than 2



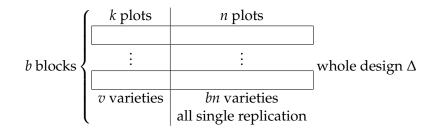
A general block design with average replication less than 2



Whole design  $\Delta$  has v + bn varieties in b blocks of size k + n;

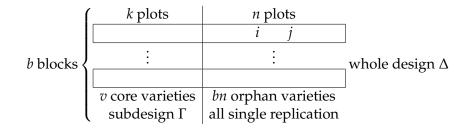


Whole design  $\Delta$  has v + bn varieties in b blocks of size k + n; the subdesign  $\Gamma$  has v core varieties in b blocks of size k;

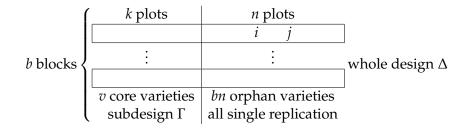


Whole design  $\Delta$  has v + bn varieties in b blocks of size k + n; the subdesign  $\Gamma$  has v core varieties in b blocks of size k; call the remaining varieties orphans.

## Pairwise variance: two orphans in the same block

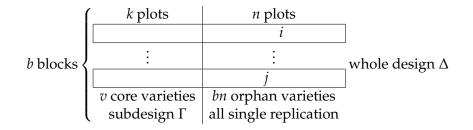


## Pairwise variance: two orphans in the same block

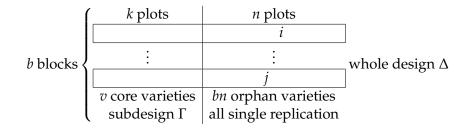


$$\operatorname{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = 2\sigma^2.$$

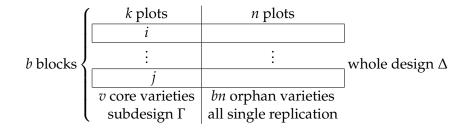
## Pairwise variance: two orphans in different blocks

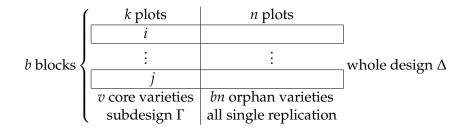


## Pairwise variance: two orphans in different blocks



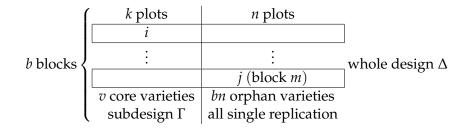
$$\operatorname{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = 2\sigma^2 + \operatorname{Var}_{\Gamma}(\hat{\beta}_i - \hat{\beta}_j).$$



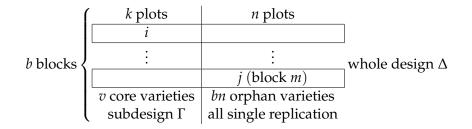


$$\operatorname{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = \operatorname{Var}_{\Gamma}(\hat{\tau}_i - \hat{\tau}_j).$$

Pairwise variance: one core variety and one orphan



Pairwise variance: one core variety and one orphan



$$\operatorname{Var}_{\Delta}(\hat{\tau}_i - \hat{\tau}_j) = \sigma^2 + \operatorname{Var}_{\Gamma}(\hat{\tau}_i + \hat{\beta}_m).$$

# Sum of the pairwise variances

Theorem (cf Herzberg and Jarrett, 2007) The sum of the variances of treatment differences in  $\Delta$ 

$$= constant + V_1 + nV_3 + n^2V_2,$$

where

- $V_1 = the sum of the variances of treatment differences in \Gamma$
- $V_2 = the sum of the variances of block differences in \Gamma$
- $V_3$  = the sum of the variances of sums of one treatment and one block in  $\Gamma$ .

(If  $\Gamma$  is equi-replicate then  $V_2$  and  $V_3$  are both increasing functions of  $V_1$ .)

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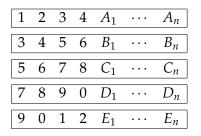
(If  $\Gamma$  is equi-replicate then  $V_2$  and  $V_3$  are both increasing functions of  $V_1$ .)

#### Consequence

For a given choice of k, make  $\Gamma$  as efficient as possible.

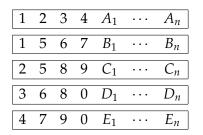
#### Consequence

If *n* or *b* is large, it may be best to make  $\Gamma$  a complete block design for *k*' controls, even if there is no interest in comparisons between new treatments and controls, or between controls.

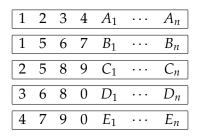


Youden and Connor (1953): "experiments in physics do not need much replication because results are not very variable" —introduced chain block designs

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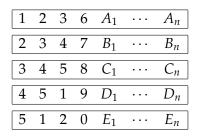
subdesign is dual of BIBD (Herzberg and Andrews, 1978)



subdesign is dual of BIBD, best subdesign for k = 4

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best subdesign for k = 3 is better for large *n* if  $b \neq 5$ 



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$$\begin{bmatrix} K_1 & K_2 & 1 & 2 & A_1 & \cdots & A_n \\ \hline K_1 & K_2 & 3 & 4 & B_1 & \cdots & B_n \\ \hline K_1 & K_2 & 5 & 6 & C_1 & \cdots & C_n \\ \hline K_1 & K_2 & 7 & 8 & D_1 & \cdots & D_n \\ \hline \hline K_1 & K_2 & 9 & 0 & E_1 & \cdots & E_n \\ \end{bmatrix}$$

better for large *n* if b > 13 even if there is no interest in controls