

# Rectangular experiments: restricted randomization or row-column designs?

R. A. Bailey



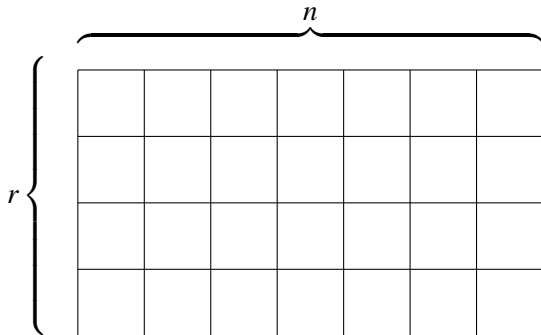
`r.a.bailey@qmul.ac.uk`

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# The problem

An agricultural experiment to compare  $n$  treatments.

The experimental area has  $r$  rows and  $n$  columns.



Use a randomized complete-block design with rows as blocks.

(In each row, choose one of the  $n!$  orders with equal probability.)

What should we do if the randomization produces a plan with one treatment always at one side of the rectangle?

# Example

Federer (1955 book): guayule trees

<i>B</i>	<i>D</i>	<i>G</i>	<i>A</i>	<i>F</i>	<i>C</i>	<i>E</i>
<i>A</i>	<i>G</i>	<i>C</i>	<i>D</i>	<i>F</i>	<i>B</i>	<i>E</i>
<i>G</i>	<i>E</i>	<i>D</i>	<i>F</i>	<i>B</i>	<i>C</i>	<i>A</i>
<i>B</i>	<i>A</i>	<i>C</i>	<i>F</i>	<i>G</i>	<i>E</i>	<i>D</i>
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<i>B</i>	<i>A</i>	<i>C</i>	<i>F</i>	<i>G</i>	<i>E</i>	<i>D</i>
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Solution: Use a carefully chosen Latinized design;  
REML/ANOVA estimates of variance components

## Assumed model

$Y_\alpha$  is the response on plot  $\alpha$ .

$E(Y_\alpha) = \theta_i$  where  $i$  is the treatment on  $\alpha$ .

$$\begin{aligned}\text{Var}(Y_\alpha) &= \sigma^2 && \text{for all } \alpha \\ \text{Cov}(Y_\alpha, Y_\beta) &= \begin{cases} \rho\sigma^2 & \text{if } \alpha \neq \beta \text{ in same row} \\ \tau\sigma^2 & \text{if } \alpha \neq \beta \text{ in same column} \\ 0 & \text{if } \alpha \neq \beta \text{ otherwise} \end{cases}\end{aligned}$$

with  $0 \leq \rho \leq 1$  and  $0 \leq \tau \leq 1$ .

## Pairwise variance in Example

<i>B</i>	<i>D</i>	<i>G</i>	<i>A</i>	<i>F</i>	<i>C</i>	<i>E</i>
<i>A</i>	<i>G</i>	<i>C</i>	<i>D</i>	<i>F</i>	<i>B</i>	<i>E</i>
<i>G</i>	<i>E</i>	<i>D</i>	<i>F</i>	<i>B</i>	<i>C</i>	<i>A</i>
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From 
$$V_{BG} = \frac{2\sigma^2}{5} \left[ 1 - \rho - \frac{4}{5}\tau \right]$$

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<i>B</i>	<i>A</i>	<i>C</i>	<i>F</i>	<i>G</i>	<i>E</i>	<i>D</i>
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From  $V_{BG} = \frac{2\sigma^2}{5} \left[ 1 - \rho - \frac{4}{5}\tau \right]$  to  $V_{EF} = \frac{2\sigma^2}{5} [1 - \rho + \tau]$

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with average  $V = \frac{2\sigma^2}{5} \left[ 1 - \rho - \frac{1}{15}\tau \right].$

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- ▶ Simple to construct.
- ▶ Simple to randomize.
- ▶ Simple to analyse.
- ▶ Some treatment comparisons in some experiments will have a specially low or specially high variance, but the estimated variance is unbiased when averaged over all comparisons and all possible randomized plans.

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- ▶ The **actual** variance of treatment comparisons is **lower**,  
but the **estimate** of that variance is **higher**.

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$$V = \frac{2\sigma^2}{r} \left[ (1 - \rho) - \frac{(r-1)\tau}{n-1} \right]$$

and

$$E(\hat{V}) = \frac{2E(M)}{r} = \frac{2\sigma^2}{r} \left[ (1 - \rho) + \frac{\tau}{n-1} \right]$$

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- ▶ Genuine treatment differences may not be detected.



## Use a Latinized design, but analyse as usual

Deliberately construct a design in which no treatment occurs more than once in any column.

Easy to do this directly, eg

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<i>G</i>	<i>A</i>	<i>B</i>	<i>C</i>	<i>D</i>	<i>E</i>	<i>F</i>
<i>B</i>	<i>C</i>	<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>	<i>A</i>
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<i>B</i>	<i>C</i>	<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>	<i>A</i>
<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>	<i>A</i>	<i>B</i>	<i>C</i>

Randomize rows, columns, treatments.

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<i>B</i>	<i>C</i>	<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>	<i>A</i>
<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>	<i>A</i>	<i>B</i>	<i>C</i>

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Same bias in estimator of variance as for simple restricted randomization.

# Super-valid restricted randomization

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- ▶ The estimator of variance is unbiased when averaged over all comparisons **in this one experiment**.

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- ▶ Analyse as usual.
- ▶ **Same** average variance as in randomized complete-block design, but with **smaller range**.
- ▶ The estimator of variance is unbiased when averaged over all comparisons **in this one experiment**.
- ▶ There is no separate estimate of  $\rho$  or  $\tau$ , so treatments must be randomized and a single standard error given for all differences.

## A design from the tables

<i>A</i>	<i>B</i>	<i>C</i>	<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>
<i>D</i>	<i>E</i>	<i>F</i>	<i>C</i>	<i>A</i>	<i>B</i>	<i>G</i>
<i>A</i>	<i>G</i>	<i>F</i>	<i>B</i>	<i>C</i>	<i>E</i>	<i>D</i>
<i>D</i>	<i>B</i>	<i>G</i>	<i>F</i>	<i>C</i>	<i>A</i>	<i>E</i>
<i>G</i>	<i>E</i>	<i>C</i>	<i>B</i>	<i>D</i>	<i>A</i>	<i>F</i>

1. In every pair of rows, there is exactly one column in which the two treatments are the same.

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<i>D</i>	<i>B</i>	<i>G</i>	<i>F</i>	<i>C</i>	<i>A</i>	<i>E</i>
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<i>A</i>	<i>G</i>	<i>F</i>	<i>B</i>	<i>C</i>	<i>E</i>	<i>D</i>
<i>D</i>	<i>B</i>	<i>G</i>	<i>F</i>	<i>C</i>	<i>A</i>	<i>E</i>
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3. If  $m_i =$  the number of columns in which treatment  $i$  occurs twice, then  $m_i - m_j \in \{-1, 0, 1\}$  for all other treatments  $j$ .

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<i>A</i>	<i>G</i>	<i>F</i>	<i>B</i>	<i>C</i>	<i>E</i>	<i>D</i>
<i>D</i>	<i>B</i>	<i>G</i>	<i>F</i>	<i>C</i>	<i>A</i>	<i>E</i>
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4. Subject to conditions (1)–(3), the spread of the variances of the estimators of simple treatment differences is as small as possible.



## Pairwise variances in the example

<i>A</i>	<i>B</i>	<i>C</i>	<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>
<i>D</i>	<i>E</i>	<i>F</i>	<i>C</i>	<i>A</i>	<i>B</i>	<i>G</i>
<i>A</i>	<i>G</i>	<i>F</i>	<i>B</i>	<i>C</i>	<i>E</i>	<i>D</i>
<i>D</i>	<i>B</i>	<i>G</i>	<i>F</i>	<i>C</i>	<i>A</i>	<i>E</i>
<i>G</i>	<i>E</i>	<i>C</i>	<i>B</i>	<i>D</i>	<i>A</i>	<i>F</i>

Minimum  $V_{AD} = \frac{2\sigma^2}{5} \left[ 1 - \rho - \frac{2}{5}\tau \right]$

Maximum  $V_{AB} = \frac{2\sigma^2}{5} \left[ 1 - \rho + \frac{2}{5}\tau \right]$

Average  $V = \frac{2\sigma^2}{5} (1 - \rho)$

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<i>A</i>	<i>G</i>	<i>F</i>	<i>B</i>	<i>C</i>	<i>E</i>	<i>D</i>
<i>D</i>	<i>B</i>	<i>G</i>	<i>F</i>	<i>C</i>	<i>A</i>	<i>E</i>
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Minimum  $V_{AD} = \frac{2\sigma^2}{5} \left[ 1 - \rho - \frac{2}{5}\tau \right] \quad \dots - \frac{4}{5}\tau$

Maximum  $V_{AB} = \frac{2\sigma^2}{5} \left[ 1 - \rho + \frac{2}{5}\tau \right] \quad \dots + \tau$

Average  $V = \frac{2\sigma^2}{5} (1 - \rho) \quad \dots - \frac{1}{15}\tau$

one layout,  
normal  
method

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<i>A</i>	<i>G</i>	<i>F</i>	<i>B</i>	<i>C</i>	<i>E</i>	<i>D</i>
<i>D</i>	<i>B</i>	<i>G</i>	<i>F</i>	<i>C</i>	<i>A</i>	<i>E</i>
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Minimum  $V_{AD} = \frac{2\sigma^2}{5} \left[ 1 - \rho - \frac{2}{5}\tau \right] \quad \dots - \frac{4}{5}\tau$

Maximum  $V_{AB} = \frac{2\sigma^2}{5} \left[ 1 - \rho + \frac{2}{5}\tau \right] \quad \dots + \tau$

Average  $V = \frac{2\sigma^2}{5}(1 - \rho) \quad \dots - \frac{1}{15}\tau \quad \dots - \frac{2}{3}\tau$

one layout,  
normal  
method

simple  
restricted  
randomization

# Efficient row-column designs

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- ▶ Randomize rows and columns.

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- ▶ Randomize rows and columns.
- ▶ More complicated analysis  
(should be available in software).

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- ▶ Unbiased estimator of the variance of every treatment contrast.
- ▶ There is no need to randomize treatments; the most important differences can be given the lowest variance.

## Example of a row-column design

<i>A</i>	<i>B</i>	<i>C</i>	<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>
<i>B</i>	<i>C</i>	<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>	<i>A</i>
<i>C</i>	<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>	<i>A</i>	<i>B</i>
<i>D</i>	<i>E</i>	<i>F</i>	<i>G</i>	<i>A</i>	<i>B</i>	<i>C</i>
<i>E</i>	<i>F</i>	<i>G</i>	<i>A</i>	<i>B</i>	<i>C</i>	<i>D</i>

$$V_{AB} = 1.044 \times \frac{2}{5}(1 - \rho - \tau)\sigma^2$$

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$$V_{AD} = 1.091 \times \frac{2}{5}(1 - \rho - \tau)\sigma^2$$

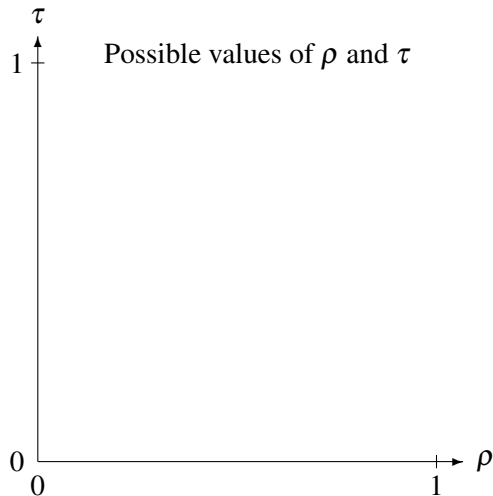
$$V = 1.075 \times \frac{2}{5}(1 - \rho - \tau)\sigma^2$$

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$$\begin{aligned} V_{AB} &= 1.044 \times \frac{2}{5}(1 - \rho - \tau)\sigma^2 && \text{normal method} \\ V_{AC} &= 1.089 \times \frac{2}{5}(1 - \rho - \tau)\sigma^2 && V = \frac{2\sigma^2}{5}(1 - \rho) \\ V_{AD} &= 1.091 \times \frac{2}{5}(1 - \rho - \tau)\sigma^2 && \text{averaged over} \\ V &= 1.075 \times \frac{2}{5}(1 - \rho - \tau)\sigma^2 && \text{randomizations} \end{aligned}$$

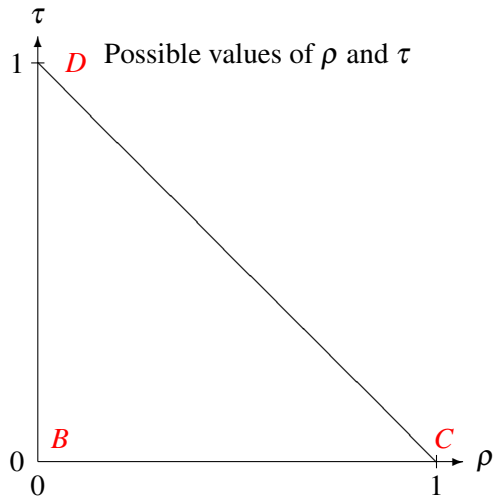
# Comparing super-valid restricted randomization and efficient row-column designs



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covariance matrix +ve definite

$$\iff \rho + \tau < 1$$



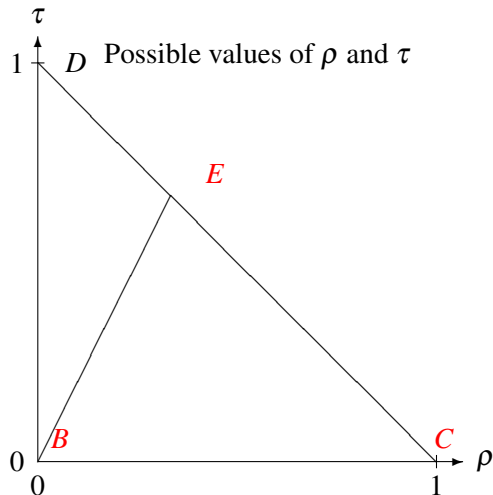
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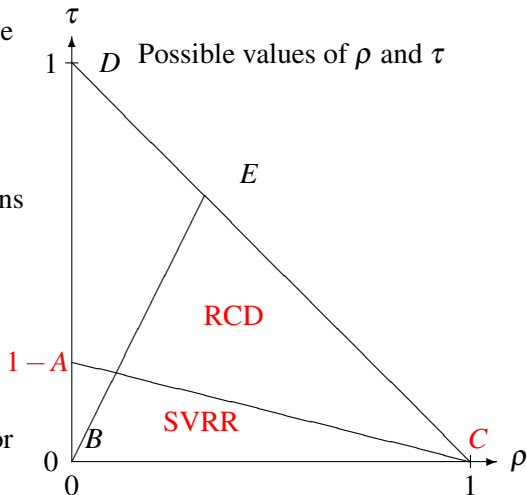
rows more variable than columns

$$\iff n\rho > r\tau$$

$V$  in RCD  $< V$  for SVRR

$$\iff (1-A)(1-\rho) < \tau$$

where  $A$  is the efficiency factor  
for the design in columns



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Choose a design with the column concurrences as equal as possible.  
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$$E(\text{MS residual from complete-block analysis}) = \sigma^2 \left[ (1 - \rho) + \frac{\tau}{n-1} \right]$$

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Hence unbiased estimators of  $\sigma^2(1 - \rho)$  and  $\sigma^2\tau$  and of

$$V = \frac{2\sigma^2}{r} \left[ (1 - \rho) - \frac{(r-1)\tau}{n-1} \right].$$

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But this estimator of  $V$  does not have a  $\chi^2$  distribution,  
so **how do we do hypothesis tests?**

Also, there are so few effective df for  $\tau$  that  
these estimates have **very poor precision**.