

Design of dose-escalation trials

R. A. Bailey



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Fisher Memorial Lecture,
Dublin, July 2008

What does a statistician do? I



What can the data tell us? Is there a simple story?

“Fisher ... pulled out—weeds.”

1952 portrait by Barrington Brown, reproduced by permission of the Fisher Memorial Trust

What does a statistician do? II



“Nature ... will best respond to a logical and carefully thought out questionnaire.”

1924 portrait, courtesy
of Joan Box

Design of the TeGenero trial

First-in-Man trial of a monoclonal antibody on healthy volunteers,
March 2006: 4 cohorts of 8 volunteers each.

Cohort	TGN1412		Placebo
	Dose mg/kg body-weight	Number of Subjects	Number of Subjects
1	0.1	6	2
2	0.5	6	2
3	2.0	6	2
4	5.0	6	2

What happened to Cohort 1 on 13 March 2006

Healthy Volunteer	Randomised to	Time of intravenous administration	Time of transfer to critical care
A	TGN1412 8.4mg	0800	2400
B	Placebo	0810	
C	TGN1412 6.8mg	0820	2350
D	TGN1412 8.8mg	0830	0030
E	TGN1412 8.2mg	0840	2040
F	TGN1412 7.2mg	0850	0050
G	TGN1412 8.2mg	0900	0100
H	Placebo	0910	

The Royal Statistical Society's Working Party on Statistical Issues in First-in-Man Studies: Membership

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What does **block** mean? **strata**? **randomize**?

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- ▶ generic issues

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(sudden adverse effects → do not dose further subjects;
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Planned analysis of the TeGenero trial

Cohort	TGN1412		Placebo
	Dose	Number	Number
1	1	6	2
2	2	6	2
3	3	6	2
4	4	6	2

If all responses are uncorrelated with variance σ^2 then
Variance (dose i – placebo) in cohort i is $(\frac{1}{6} + \frac{1}{2}) \sigma^2 = \frac{2}{3} \sigma^2$

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There have been many trials, in many topics, where, with hindsight, cohort effects swamp treatment effects.

The Experimental Medicines Group of the Association of the British Pharmaceutical Industry (ABPI) says that trials should always be designed on the assumption that there will be cohort effects.

Analysis of the TeGenero trial with cohort effects

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$$\text{Variance (dose } i - \text{ placebo) in cohort } i = \left(\frac{1}{6} + \frac{1}{2} \right) \sigma^2 = \frac{2}{3} \sigma^2.$$

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Estimator of (dose i – dose j) =

$$\begin{aligned} & [\text{estimator of (dose } i - \text{ placebo) in cohort } i] - \\ & [\text{estimator of (dose } j - \text{ placebo) in cohort } j] \end{aligned}$$

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$$\text{So variance (dose } i - \text{ dose } j) = \left(\frac{2}{3} + \frac{2}{3} \right) \sigma^2 = \frac{4}{3} \sigma^2.$$

Senn's proposed design

Cohort	TGN1412		Placebo
	Dose	Number	Number
1	1	4	4
2	2	4	4
3	3	4	4
4	4	4	4

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So variance (dose i – dose j) = $\left(\frac{1}{2} + \frac{1}{2}\right) \sigma^2 = \sigma^2$

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The TeGenero design is **inadmissible** because everything can be estimated, from the same resources, with smaller variance, by another design.

Dose-escalation trials: standard designs

There are n doses, with dose $1 < \text{dose } 2 < \dots < \text{dose } n$.

0 denotes the placebo.

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Put s_{ki} = number of subjects who get dose i in cohort k . Then

$$s_{ki} > 0 \quad \text{if} \quad i = k$$

$$s_{ki} = 0 \quad \text{if} \quad i > k.$$

Scaled variance

Assume that the expectation of the response of a subject who gets dose i in cohort k is $\tau_i + \beta_k$, and that responses are uncorrelated with common variance σ^2 .

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If doses could be equally replicated within each cohort, then each pairwise variance would be

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so define the **scaled variance** v_{ij} to be

$$\frac{\text{Variance (dose } i - \text{dose } j) \times \text{number of observations}}{2(n+1)\sigma^2}.$$

Textbook design

Aim:

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Example: $n = 4, m = 10$

Dose	0	1	2	3	4
Cohort 1	2	8	0	0	0
Cohort 2	2	0	8	0	0
Cohort 3	2	0	0	8	0
Cohort 4	2	0	0	0	8

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$$v_{0i} = \frac{n+1}{2}$$

$$v_{ij} = n + 1$$

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Example: $n = 4, m = 8$

Dose	0	1	2	3	4
Cohort 1	4	4	0	0	0
Cohort 2	4	0	4	0	0
Cohort 3	4	0	0	4	0
Cohort 4	4	0	0	0	4

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$$v_{0i} = \frac{2n}{n+1} \quad v_{ij} = \frac{4n}{n+1}$$

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Principle

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Each cohort should have as many different treatments as possible.

Proposed “halving” designs

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In Cohort 1: $\frac{m}{2}$ subjects get dose 1; $\frac{m}{2}$ subjects get placebo.

In Cohort k : $\frac{m}{2}$ subjects get dose k ; remaining subjects are allocated to placebo and doses 1 to $k - 1$ according to some rule.

Strict halving design

Remaining subjects are allocated as in Cohort $k - 1$ with numbers halved.

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Example: $n = 4$, $m = 16$

Dose	0	1	2	3	4
Cohort 1	8	8	0	0	0
Cohort 2	4	4	8	0	0
Cohort 3	2	2	4	8	0
Cohort 4	1	1	2	4	8

Uniform halving design

Remaining subjects are allocated as equally as possible to treatments 0 to $k - 1$, with larger values given to make the ‘replication so far’ as equal as possible.

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Example: $n = 4$, $m = 8$

Dose	0	1	2	3	4
Cohort 1	4	4	0	0	0
Cohort 2	2	2	4	0	0
Cohort 3	1	1	2	4	0
Cohort 4	1	1	1	1	4

Catch-up halving design

Remaining subjects are allocated to treatments 0 to $k - 1$ to make the ‘replication so far’ as equal as possible, with lower doses favoured if there is a tie.

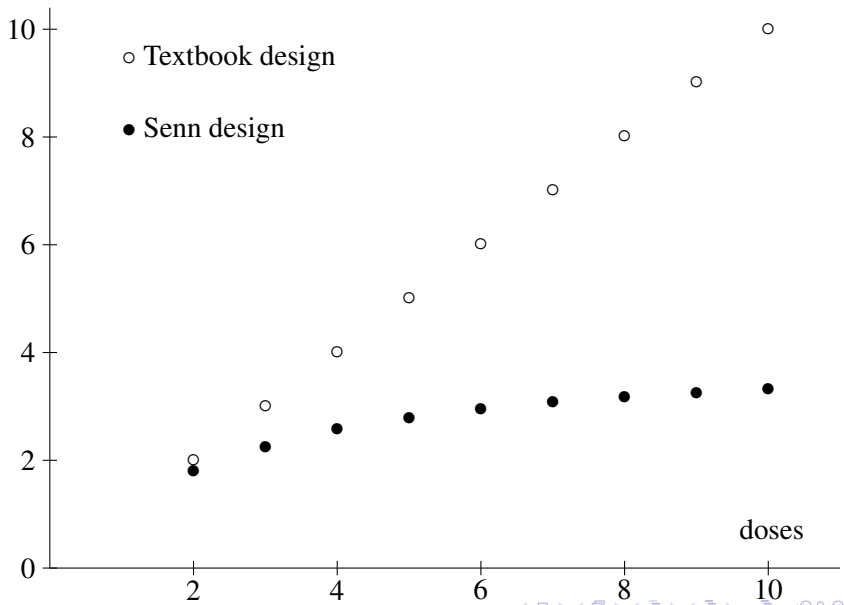
Catch-up halving design

Remaining subjects are allocated to treatments 0 to $k - 1$ to make the ‘replication so far’ as equal as possible, with lower doses favoured if there is a tie.

Example: $n = 4$, $m = 10$

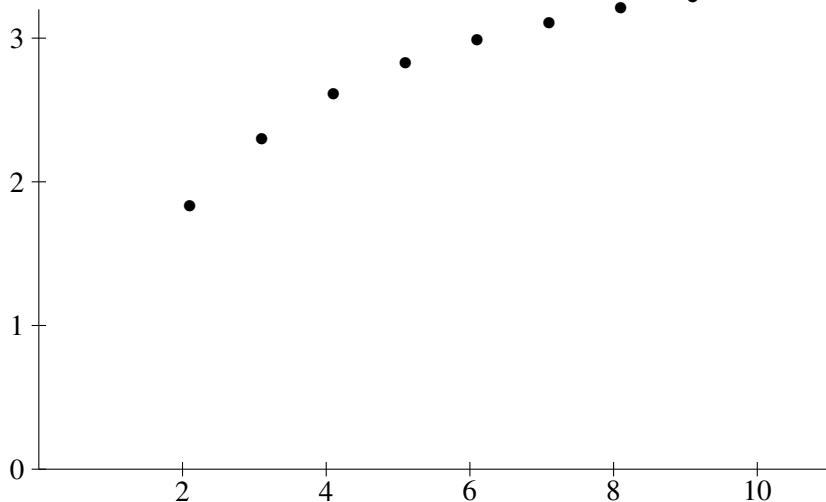
Dose	0	1	2	3	4
Cohort 1	5	5	0	0	0
Cohort 2	3	2	5	0	0
Cohort 3	1	1	3	5	0
Cohort 4	1	1	1	2	5

Average scaled pairwise variance



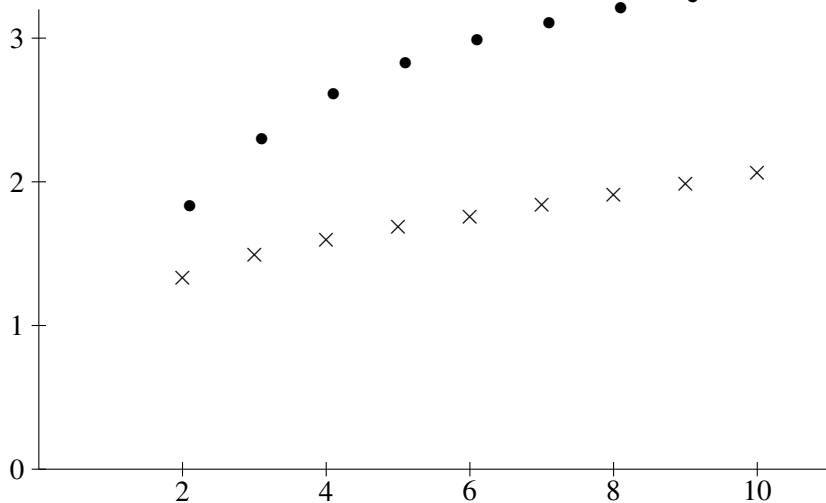
Average scaled pairwise variance: continued

- Senn design



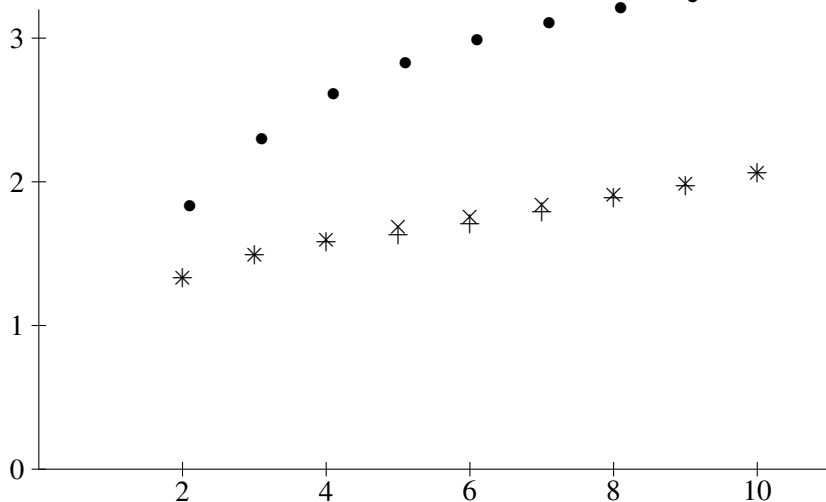
Average scaled pairwise variance: continued

- Senn design
- × strict halving design



Average scaled pairwise variance: continued

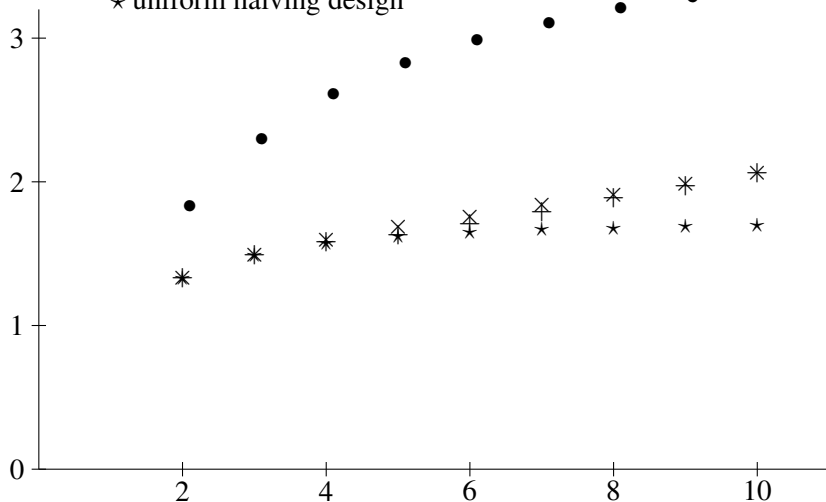
- Senn design × strict halving design + catch-up halving design



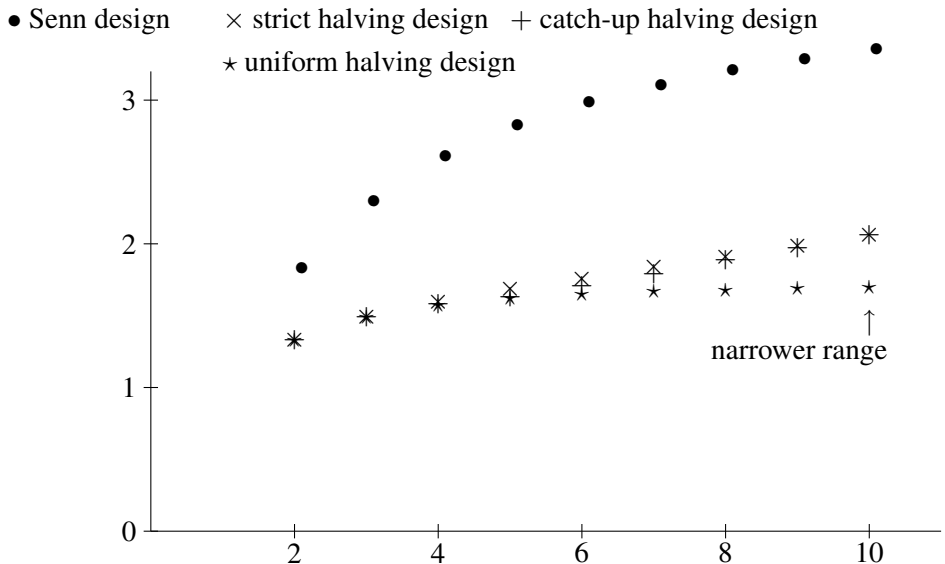
Average scaled pairwise variance: continued

● Senn design × strict halving design + catch-up halving design

★ uniform halving design



Average scaled pairwise variance: continued



Lessons from experience with block designs: II

If we transpose the matrix,
we interchange the roles of doses and cohorts,
to obtain the **dual** block design.

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In the standard designs, the highest dose has **all** of its subjects in the final cohort, so no contrast between this cohort and other cohorts can be orthogonal to that dose.

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Principle

There should be one more cohort than there are doses.

Dose-escalation trials: extended designs

There are n doses, with dose $1 < \text{dose } 2 < \dots < \text{dose } n$.

0 denotes the placebo.

There are $n + 1$ cohorts of m subjects each.

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0 denotes the placebo.

There are $n + 1$ cohorts of m subjects each.

Cohort 1 subjects may receive only dose 1 or placebo.

In Cohort i , for $2 \leq i \leq n$, some subjects receive dose i ;
no subject receives dose j if $j > i$.

In Cohort $n + 1$, any dose, or placebo, may be used.

Extended textbook design

Maintain overall equal replication in the final cohort.

$$s_{n+1,i} = \frac{m}{n+1} \quad \text{for } i = 0, \dots, n$$

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Example: $n = 4$, $m = 10$

Dose	0	1	2	3	4
Cohort 1	2	8	0	0	0
Cohort 2	2	0	8	0	0
Cohort 3	2	0	0	8	0
Cohort 4	2	0	0	0	8
Cohort 5	2	2	2	2	2

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$$v_{0i} = \frac{(n+1)(n+2)}{2(2n+1)} \quad v_{ij} = \frac{(n+1)^2}{2n+1}$$

Extended Senn design

In the final cohort,
compensate for the previous over-replication of placebo.

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Dose	0	1	2	3	4
Cohort 1	4	4	0	0	0
Cohort 2	4	0	4	0	0
Cohort 3	4	0	0	4	0
Cohort 4	4	0	0	0	4
Cohort 5	0	2	2	2	2

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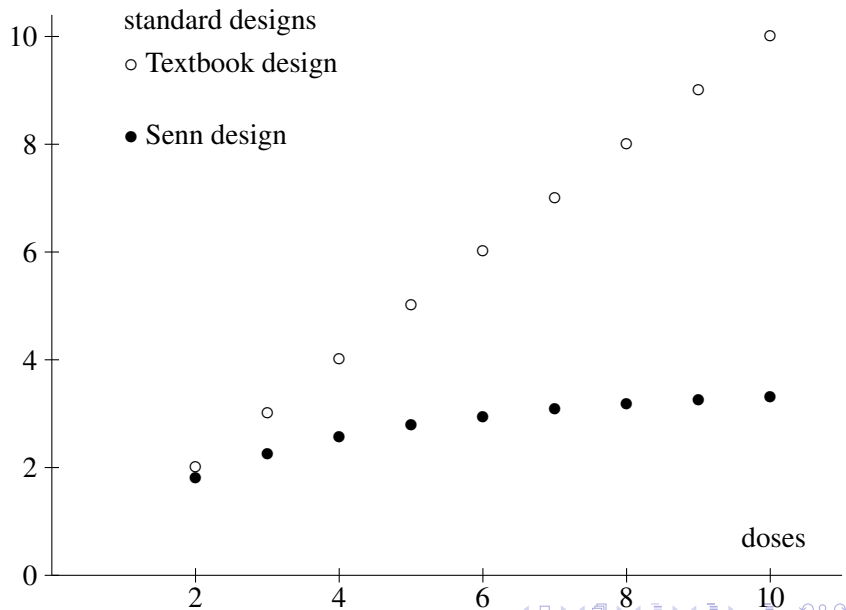
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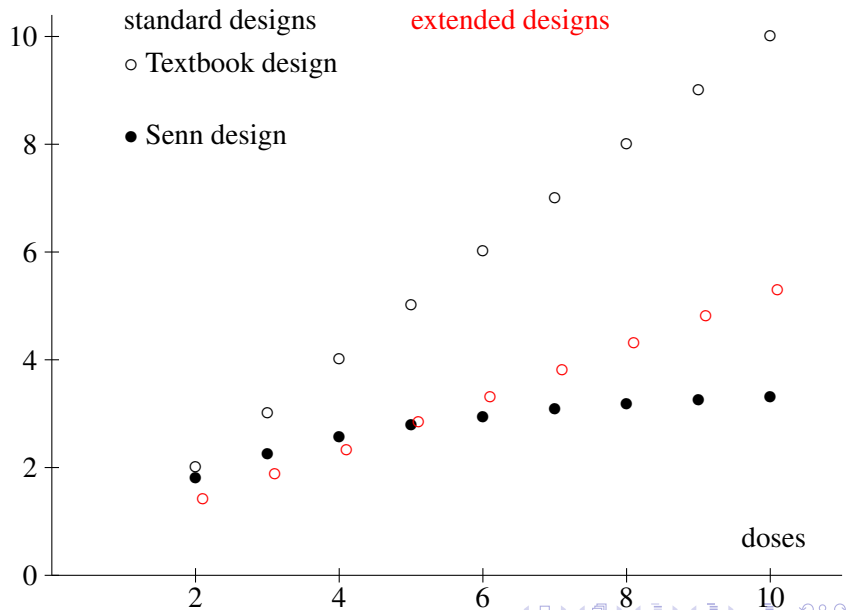
Dose	0	1	2	3	4
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Cohort 2	4	0	4	0	0
Cohort 3	4	0	0	4	0
Cohort 4	4	0	0	0	4
Cohort 5	0	2	2	2	2

$$v_{0i} = \frac{2(n^2 + 4)}{n(n + 4)} \quad v_{ij} = \frac{4n}{n + 4}$$

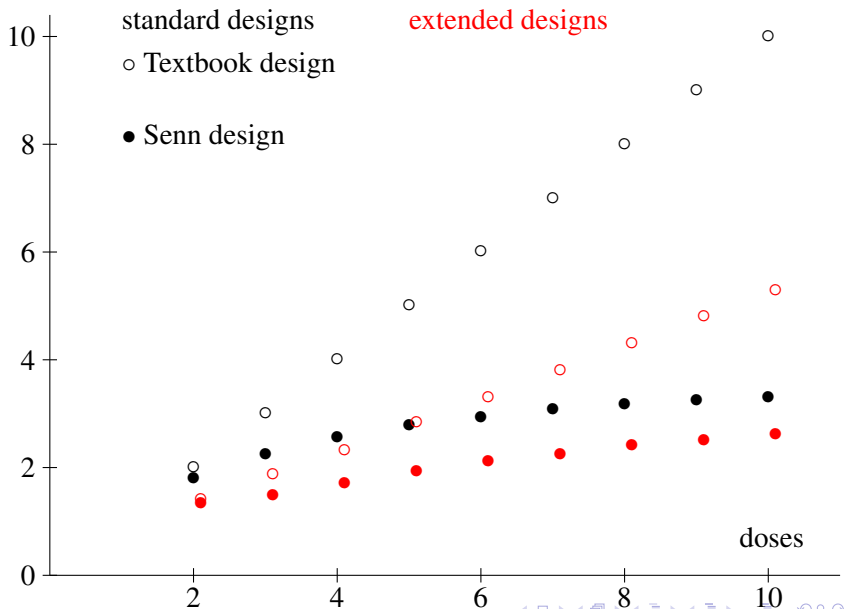
Average scaled pairwise variance (again)



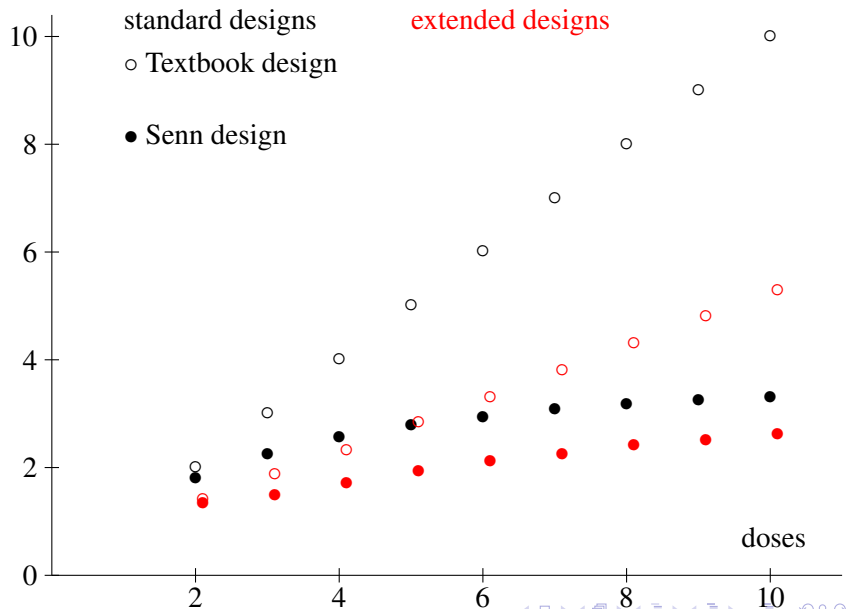
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Extended halving designs

In all of the standard halving designs,
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The extra cohort should try to redress this.

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the largest pairwise variances are those involving the highest dose.
The extra cohort should try to redress this.

Is it better to aim for equal replication in the final cohort,
or equal replication overall?

Uniform extension

The replications in the extra cohort as are equal as possible, with larger values given to those doses with lower replication so far.

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Example: $n = 4$, $m = 8$

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Equireplicate extension

The replications in the extra cohort make the overall replications as equal as possible, with any inequalities favouring the higher doses.

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About half the subjects in the final cohort are equally split between all treatments,
the remainder being allocated to make the overall replications as equal as possible, with any inequalities favouring the higher doses.

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	1	1	1	1	1

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	1	1	1	1	1
					1
				1	1

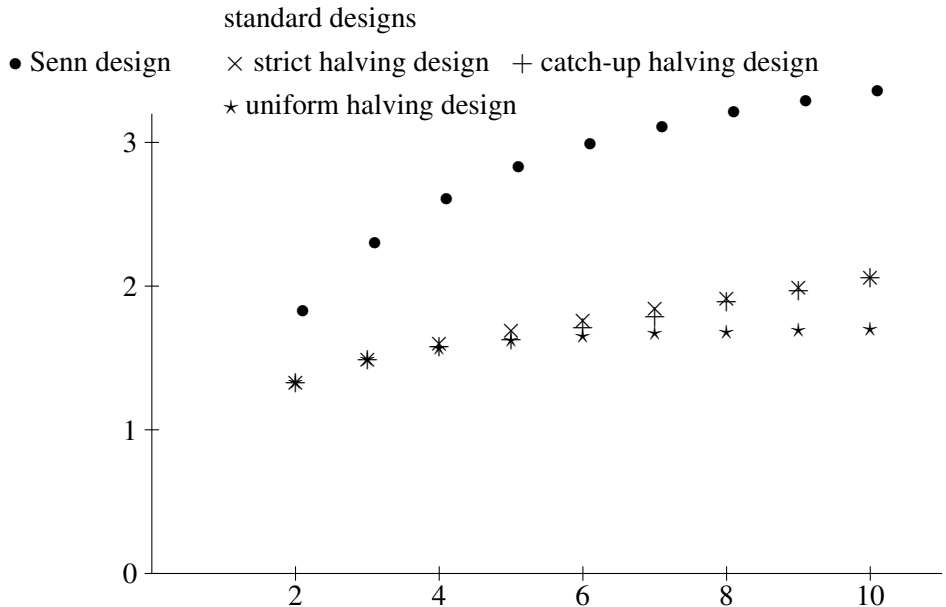
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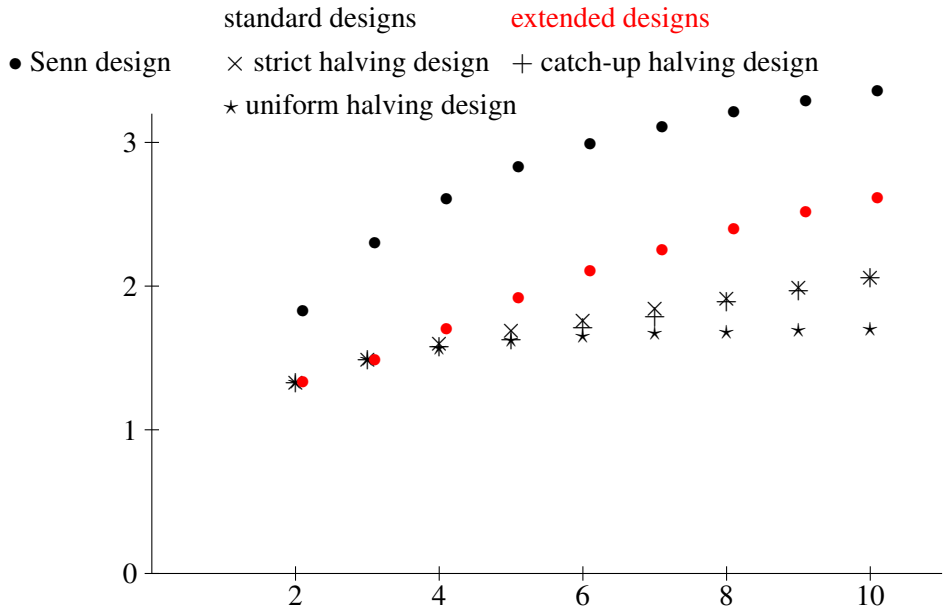
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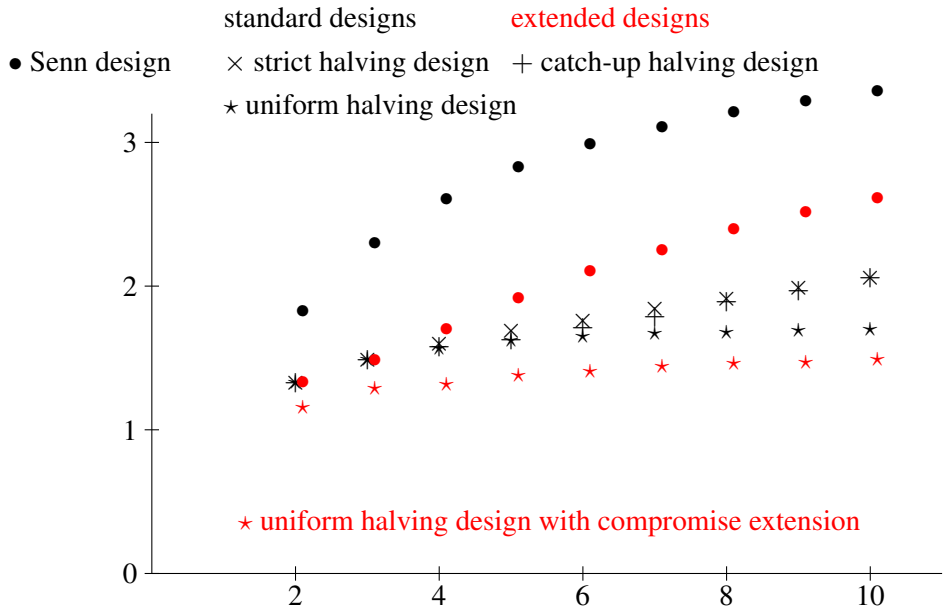
Average scaled pairwise variance: continued (again)



Average scaled pairwise variance: continued (again)



Average scaled pairwise variance: continued (again)



Two designs for 4 doses using 40 subjects

Numbers of subjects							Actual pairwise variances/ σ^2				
St T	Dose	0	1	2	3	4		1	2	3	4
	Cohort 1	2	8	0	0	0	0	0.625	0.625	0.625	0.625
	Cohort 2	2	0	8	0	0	1		1.250	1.250	1.250
	Cohort 3	2	0	0	8	0	2			1.250	1.250
	Cohort 4	2	0	0	0	8	3				1.250
Ex UH Co	Dose	0	1	2	3	4		1	2	3	4
	Cohort 1	4	4	0	0	0	0	0.222	0.285	0.348	0.370
	Cohort 2	2	2	4	0	0	1		0.285	0.348	0.370
	Cohort 3	1	1	2	4	0	2			0.330	0.378
	Cohort 4	1	1	1	1	4	3				0.375
	Cohort 5	1	1	1	2	3					

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	Cohort 3	2	0	0	8	0	2			1.250	1.250
	Cohort 4	2	0	0	0	8	3				1.250
								average 1.00			
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	Cohort 2	2	2	4	0	0	1		0.285	0.348	0.370
	Cohort 3	1	1	2	4	0	2			0.330	0.378
	Cohort 4	1	1	1	1	4	3				0.375
	Cohort 5	1	1	1	2	3		average 0.33			

Random cohort effects

Now assume that the expectation of the response of a subject who gets dose i in cohort k is τ_i , and that cohort effects are uncorrelated random variables with common variance σ_C^2 .

$$\text{Put } \mathbf{C}_{\alpha\beta} = \begin{cases} 1 & \text{if subjects } \alpha \text{ and } \beta \text{ are in the same cohort} \\ 0 & \text{otherwise.} \end{cases}$$

Then the variance-covariance matrix of the responses is

$$\sigma^2 \mathbf{I} + \sigma_C^2 \mathbf{C}$$

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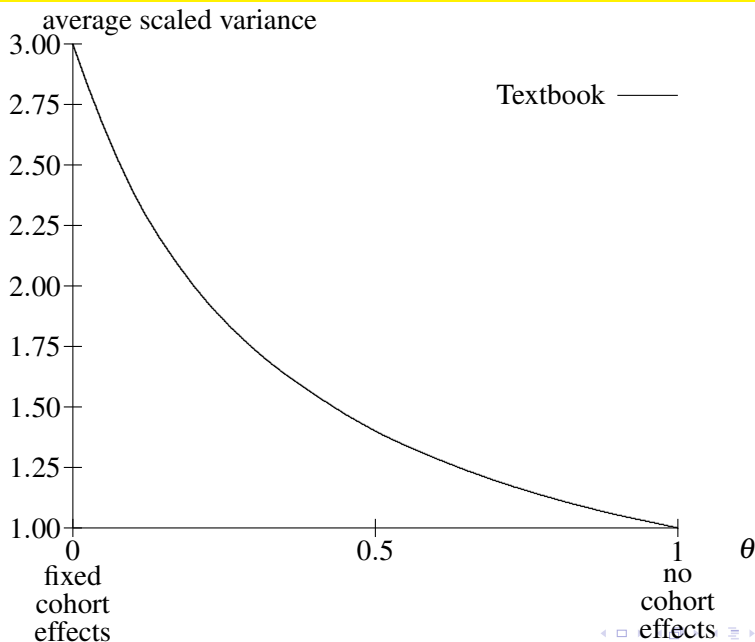
$$\sigma^2 \mathbf{I} + \sigma_C^2 \mathbf{C} = \underbrace{\sigma^2 \left(\mathbf{I} - \frac{1}{m} \mathbf{C} \right)}_{\text{within cohorts}} + \underbrace{\sigma^2 \theta^{-1} \frac{1}{m} \mathbf{C}}_{\text{between cohorts}}$$

where $\sigma^2 + m\sigma_C^2 = \theta^{-1}\sigma^2$,

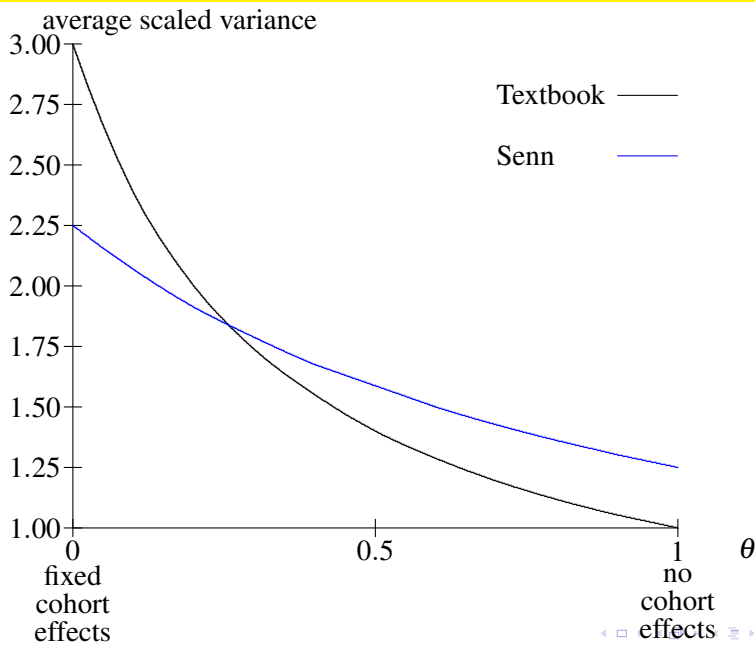
so $\theta \in [0, 1]$ with $\theta = 0$ if cohort effects are fixed

$\theta = 1$ if cohort effects are zero.

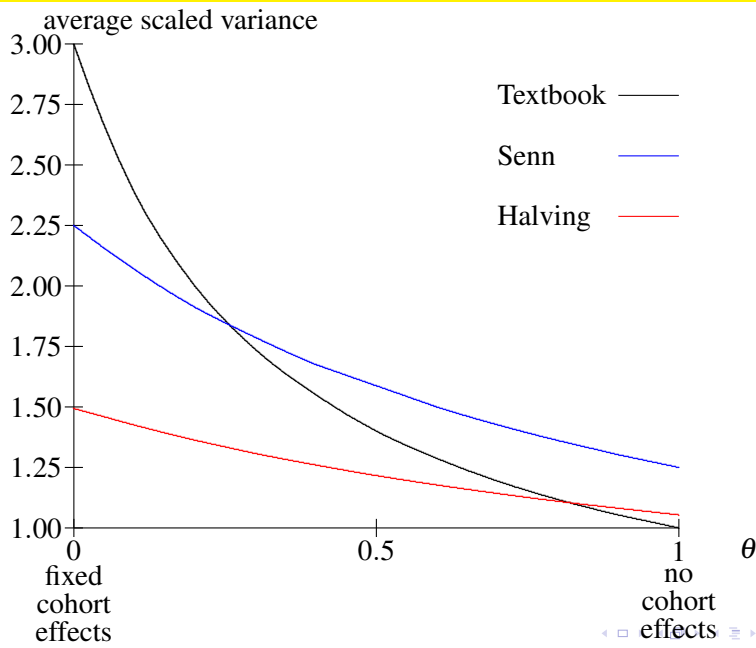
Average scaled variance for 3 doses in 3 cohorts



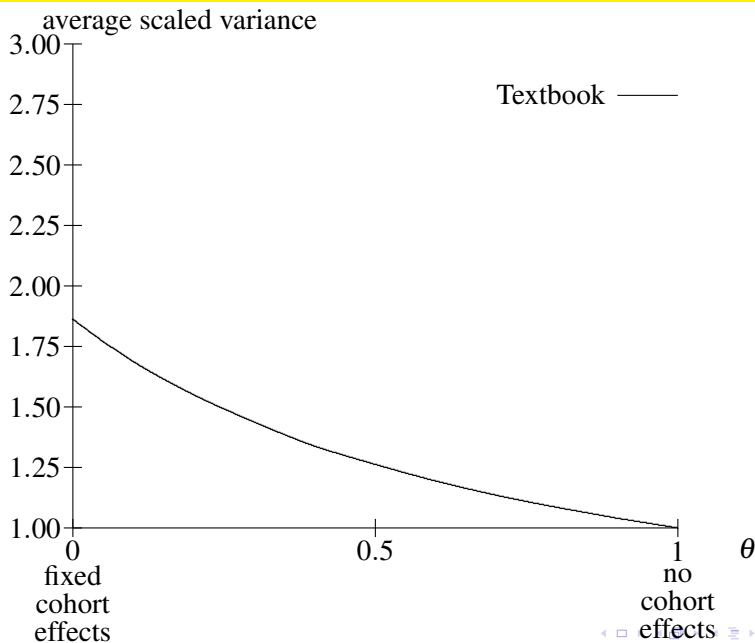
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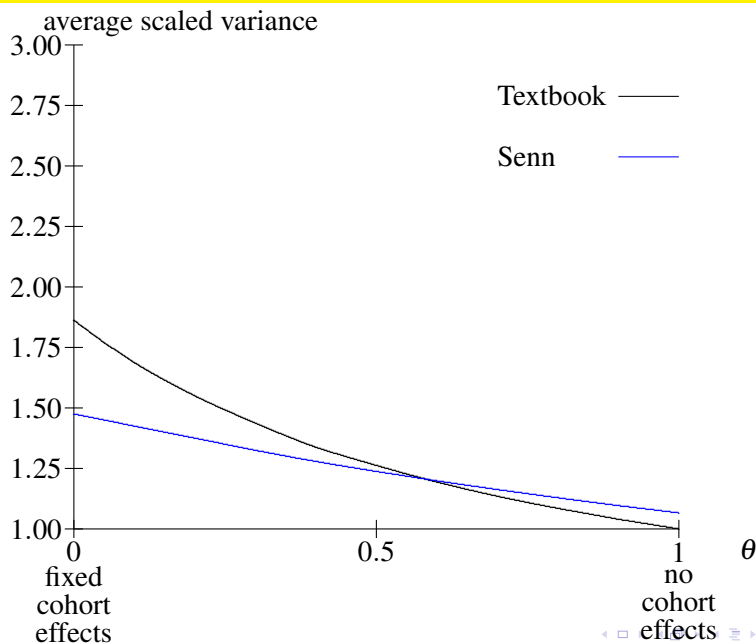
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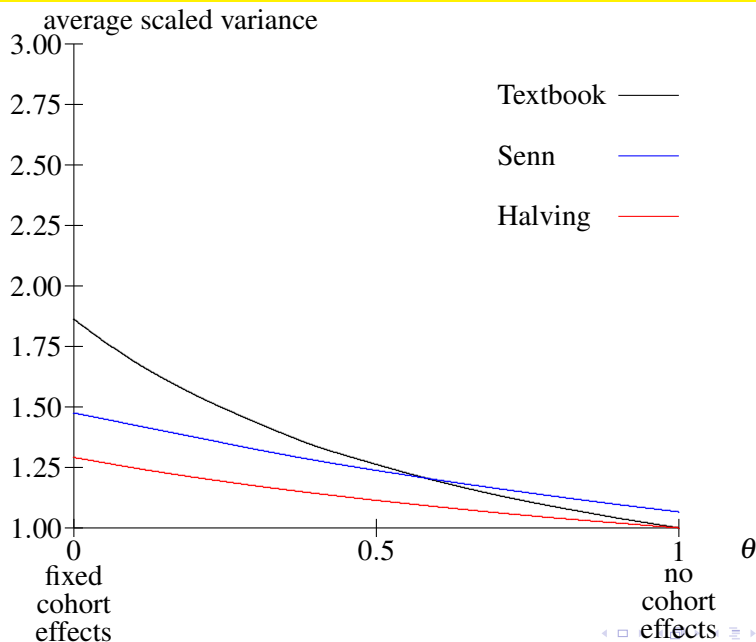
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Among the extended designs examined, the best are the uniform halving designs with the compromise extension.

Both types can be described by the following simple rule:

Principle

*In each cohort,
half of the subjects should be distributed (approximately) equally
among all the treatments that have been used in any previous cohort;
the remaining subjects should be used to make the replication so far
as equal as possible by compensating for previous under-replication.*

Advantages of the halving designs

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- ▶ If cohort effects are small and random, the variance is very little more than for the textbook design.
- ▶ Blinding is more effective than in textbook designs.

R. A. Fisher Memorial Lecture



Applying experience from agricultural field trials to dose-escalation trials in humans.

I hope he would have approved.