

## Probability III – 2008/09

### Solutions to Exercise Sheet 7

In lectures, we proved several statements about the Birth Process (BP). They can be briefly summarized as follows.

**Theorem 1.** Suppose that  $X(t)$  is a birth process with  $X(0) = 0$ . Set  $p_n(t) \stackrel{\text{def}}{=} P\{X(t) = n | X(0) = 0\}$ . Then the functions  $p_n(t)$  satisfy the following equations:

$$\begin{cases} p'_n(t) = -\lambda_n p_n(t) + \lambda_{n-1} p_{n-1}(t) & n \geq 0 \\ p_0(0) = 1 \\ p_n(0) = 0 & \text{if } n > 0. \end{cases} \quad (1)$$

**Theorem 2.** Equations (1) have a unique solution which can be found recursively using the following formulae:

$$\begin{cases} p_0(t) = e^{-\lambda_0 t} \\ p_n(t) = \lambda_{n-1} e^{-\lambda_n t} \int_0^t e^{\lambda_n s} p_{n-1}(s) ds & n > 0 \end{cases} \quad (2)$$

Marks

[60]

1. Let  $X(t)$  be a Birth Process with parameters  $\lambda_0, \lambda_1, \dots, \lambda_n, \dots$  and suppose that  $X(0) = 3$ . Set  $p_n(t) \stackrel{\text{def}}{=} P\{X(t) = n | X(0) = 3\}$ 
  - (a) Using the method explained in lectures derive equations for  $p_3(t)$  and for  $p_4(t)$ .

**Solution**

Since  $X(0) = 3$ , state 3 plays in our case the same role as 0 in the case when  $X(0) = 0$ . The only difference is that now the parameters of the Birth Process are  $\lambda_3, \lambda_4, \dots$  (remark that  $\lambda_0, \lambda_1, \lambda_2$  are unimportant because  $X(t) \geq 3$ ). We can thus claim by analogy that

$$\begin{cases} p'_n(t) = -\lambda_n p_n(t) + \lambda_{n-1} p_{n-1}(t) & n \geq 3 \\ p_3(0) = 1 \\ p_n(0) = 0 & \text{if } n > 3. \end{cases} \quad (3)$$

In particular, for  $n = 3$  and taking into account that  $p_2(t) \equiv 0$  we claim that

$$p'_3(t) = -\lambda_3 p_3(t), \quad \text{with } p_3(0) = 1. \quad (4)$$

If  $n = 4$  then the claim is that

$$p'_4(t) = -\lambda_4 p_4(t) + \lambda_3 p_3(t) \quad \text{with } p_4(0) = 0 \quad (5)$$

The proof is exactly the same as the one explained in lectures (with natural replacements of  $\lambda_0$  by  $\lambda_3$  etc. Nevertheless you are required to write it down. So here it is.

**Proof of (4).** For  $h > 0$ , the following equality follows from the monotonicity of the Birth Process (if  $X(0) = 3$  and  $X(t+h) = 3$ , then also  $X(t) = 3$ )

$$\begin{aligned} P\{X(t+h) = 3|X(0) = 3\} &= P\{X(t) = 3, X(t+h) = 3|X(0) = 3\} \\ &= P\{X(t) = 3|X(0) = 3\}P\{X(t+h) = 3|X(t) = 3, X(0) = 3\}. \end{aligned} \quad (6)$$

By the Markov property

$$P\{X(t+h) = 3|X(t) = 3, X(0) = 3\} = P\{X(t+h) = 3|X(t) = 3\}.$$

Equation (6) can now be written as

$$p_3(t+h) = p_3(t)P\{X(t+h) = 3|X(t) = 3\} = p_3(t)(1 - \lambda_3 h + o(h)), \quad (7)$$

where the quality  $P\{X(t+h) = 3|X(t) = 3\} = 1 - \lambda_3 h + o(h)$  is due to the definition of the Birth process. Equation (7) can be written as

$$p_3(t+h) = p_3(t) - \lambda_3 h p_3(t) + o(h) \quad \text{remember that } p_3(t)o(h) = o(h).$$

Hence

$$\frac{p_3(t+h) - p_3(t)}{h} = -\lambda_3 p_3(t) + \frac{o(h)}{h}.$$

Letting  $h \rightarrow 0$  implies

$$p_3'(t) = -\lambda_3 p_3(t),$$

and this finishes the proof of (4). Finally, the initial condition  $p_3(0) = 1$  follows from  $p_3(0) = P\{X(0) = 3|X(0) = 3\} = 1$  which is clearly true.

**Proof of (5).** For  $h > 0$ , the following equality follows from the monotonicity property of the BP and the total probability formula:

$$\begin{aligned} P\{X(t+h) = 4|X(0) = 3\} &= \\ P\{X(t) = 3|X(0) = 3\}P\{X(t+h) = 4|X(t) = 3, X(0) = 3\} &+ \\ P\{X(t) = 4|X(0) = 3\}P\{X(t+h) = 4|X(t) = 4, X(0) = 3\}. \end{aligned} \quad (8)$$

By the Markov property

$$\begin{aligned} P\{X(t+h) = 4|X(t) = 3, X(0) = 3\} &= P\{X(t+h) = 4|X(t) = 3\}, \\ P\{X(t+h) = 4|X(t) = 4, X(0) = 3\} &= P\{X(t+h) = 4|X(t) = 4\} \end{aligned}$$

Equation (8) can now be written as

$$p_4(t+h) = p_3(t)P\{X(t+h) = 4|X(t) = 3\} + p_4(t)P\{X(t+h) = 4|X(t) = 4\}. \quad (9)$$

Since by the definition of the Birth Process  $P\{X(t+h) = 4|X(t) = 3\} = \lambda_3 h + o(h)$  and  $P\{X(t+h) = 4|X(t) = 4\} = 1 - \lambda_4 h + o(h)$  we obtain

$$p_4(t+h) = p_4(t)(1 - \lambda_4 h + o(h)) + p_3(t)(\lambda_3 h + o(h)).$$

Hence

$$\frac{p_4(t+h) - p_4(t)}{h} = -\lambda_4 p_4(t) + \lambda_3 p_3(t) + \frac{o(h)}{h}.$$

Letting  $h \rightarrow 0$  implies that

$$p_4'(t) = -\lambda_4 p_4(t) + \lambda_3 p_3(t),$$

and this finishes the proof of (5). The initial condition  $p_4(0) = 0$  follows from  $p_4(0) = P\{X(0) = 4|X(0) = 3\} = 0$ .

(b) State the equation for  $p_n(t)$ .

**Solution** It is the same equation as before, namely

$$p_n'(t) = -\lambda_n p_n(t) + \lambda_{n-1} p_{n-1}(t). \quad (10)$$

The difference is that now  $n \geq 4$  and that  $p_3(t) = e^{-\lambda_3 t}$  (obviously,  $p_i(t) \equiv 0$ ,  $0 \leq i \leq 2$ ).

(c) Derive formulae similar to (2) for  $p_3(t)$  and for  $p_n(t)$ ,  $n \geq 4$ .

**Solution** For  $n = 3$ , it follows from (4) that  $p_3(t) = ce^{-\lambda_3 t}$ . Since  $p_3(0) = 1$ , we have  $ce^{-\lambda_3 \times 0} = 1$ . Hence  $c = 1$  and  $p_3(t) = e^{-\lambda_3 t}$ .

If  $n \geq 4$  then multiplying equation (10) by  $e^{\lambda_n t}$  and writing it as

$$e^{\lambda_n t} p_n'(t) + e^{\lambda_n t} \lambda_n p_n(t) = \lambda_{n-1} e^{\lambda_n t} p_{n-1}(t), \quad n \geq 0$$

we obtain

$$(e^{\lambda_n t} p_n(t))' = \lambda_{n-1} e^{\lambda_n t} p_{n-1}(t), \quad n \geq 3.$$

(make sure that you understand this calculation!). Hence

$$\int_0^t (e^{\lambda_n s} p_n(s))' ds = \lambda_{n-1} \int_0^t e^{\lambda_n s} p_{n-1}(s) ds, \quad n \geq 3$$

and we obtain that

$$e^{\lambda_n t} p_n(t) - p_n(0) = \lambda_{n-1} \int_0^t e^{\lambda_n s} p_{n-1}(s) ds, \quad n \geq 3.$$

Thus

$$p_n(t) = e^{-\lambda_n t} p_n(0) + e^{-\lambda_n t} \lambda_{n-1} \int_0^t e^{\lambda_n s} p_{n-1}(s) ds, \quad n \geq 3.$$

If  $n = 3$  then  $p_3(0) = P\{X(0) = 3|X(0) = 3\} = 1$ ,  $p_2(s) \equiv 0$  and the last formula implies that  $p_3(t) = e^{-\lambda_3 t}$  (which means that we got one more derivation of the expression for  $p_3(t)$ ). For  $n > 3$ , we have  $p_n(0) = P\{X(0) = n|X(0) = 3\} = 0$  and

$$p_n(t) = e^{-\lambda_n t} \lambda_{n-1} \int_0^t e^{\lambda_n s} p_{n-1}(s) ds.$$

- (d) Suppose that  $\lambda_3 = 1$ ,  $\lambda_4 = 1.5$ . Find the expressions for  $p_3(t)$  and for  $p_4(t)$ .

**Solution** We have just shown (this can also be inferred from (4)) that  $p_3(t) = e^{-\lambda_3 t}$ . Next, for  $n = 4$  we have

$$p_4(t) = \lambda_3 e^{-\lambda_4 t} \int_0^t e^{\lambda_4 s} p_3(s) ds = \lambda_3 e^{-\lambda_4 t} \int_0^t e^{(\lambda_4 - \lambda_3)s} ds = \frac{\lambda_3}{\lambda_4 - \lambda_3} (e^{-\lambda_3 t} - e^{-\lambda_4 t}).$$

Hence the answer:

$$p_3(t) = e^{-t}, \quad p_4(t) = 2(e^{-t} - e^{-1.5t}).$$

- (e) Find the probability density function of the time  $W_4$  the process  $X(t)$  remains in the state 3.

**Solution** The distribution function of  $W_4$  is given by

$$F_{W_4}(y) = P\{W_4 \leq y | X(0) = 3\} = 1 - P\{W_4 > y | X(0) = 3\}.$$

But  $P\{W_4 > y | X(0) = 3\} = P\{X(y) = 3 | X(0) = 3\} = e^{-y}$  and hence  $F_{W_4}(y) = 1 - e^{-y}$  and the p. d. f. is given by  $f_{W_4}(y) = (1 - e^{-y})' = e^{-y}$  (note that  $y \geq 0$  since  $W_4 \geq 0$ ).

- (f) Hence find the mean time  $E(W_4 | X(0) = 3)$  the process  $X(t)$  spends in the state 3.

**Solution**  $E(W_4 | X(0) = 3) = \int_0^\infty y e^{-y} dy = 1$ .

- [40] 2. Consider the birth process (which we briefly discussed in lectures) with  $\lambda_n = n\lambda$  ( $\lambda > 0$ ).

- (a) Prove that equations (1) imply that  $P\{X(t) = 0 | X(0) = 0\} = 1$ .

**Proof.** Note that  $\lambda_0 = 0$ . By the definition of  $p_n(t)$  we have that  $p_0(t) = P\{X(t) = 0 | X(0) = 0\} = e^{\lambda_0 t} = e^{0 \times t} = 1$ .

- (b) If, however,  $X(0) = 1$  then

$$p_n(t) \stackrel{\text{def}}{=} P\{X(t) = n | X(0) = 1\} = e^{-\lambda t} (1 - e^{-\lambda t})^{n-1}. \quad (11)$$

Check that this is true.

Hint: since you are given the explicit expressions for  $p_n(t)$ , it suffices to show that these functions satisfy equations (1) (and relevant initial conditions).

**Solution** Let us find the derivative of  $p_n(t)$ :

$$p'_n(t) = -\lambda e^{-\lambda t} (1 - e^{-\lambda t})^{n-1} + \lambda(n-1) e^{-2\lambda t} (1 - e^{-\lambda t})^{n-2}.$$

Hence

$$p'_n(t) = -n\lambda e^{-\lambda t} (1 - e^{-\lambda t})^{n-1} + (n-1)\lambda e^{-\lambda t} (1 - e^{-\lambda t})^{n-1} + \lambda(n-1) e^{-2\lambda t} (1 - e^{-\lambda t})^{n-2}.$$

In other words,

$$p'_n(t) = -\lambda_n p_n(t) + (n-1)\lambda e^{-\lambda t} (1 - e^{-\lambda t})^{n-2} (1 - e^{-\lambda t} + e^{-\lambda t})$$

and thus

$$p'_n(t) = -\lambda_n p_n(t) + \lambda_{n-1} p_{n-1}(t).$$

Since  $p_n(t)$  defined by (11) also satisfies  $p_1(0) = 1$  and  $p_n(0) = 0$  for  $n > 1$  we conclude that  $p_n(t) = P\{X(t) = n | X(0) = 1\}$ . (We use here the fact that these probabilities satisfy equations (1) with appropriate initial conditions and that there is only one sequence of functions satisfying these equations.)

- (c) The statement made in (b) means that the random variable  $X(t)$  conditioned on  $X(0) = 1$  has a geometric distribution. Hence, find  $E(X(t) | X(0) = 1)$ . Does  $E(X(t) | X(0) = 1)$  grow exponentially fast as a function of  $t$ ?

**Solution**

$$E(X(t) | X(0) = 1) = \sum_{n=1}^{\infty} n p_n(t) = \sum_{n=1}^{\infty} e^{-\lambda t} (1 - e^{-\lambda t})^{n-1} = \frac{1}{e^{-\lambda t}} = e^{\lambda t}.$$

The answer is YES, this expectation grows exponentially fast.